

**Positivity Resonance in Long-Term Married Couples:
Multimodal Characteristics and Consequences for Health and Longevity**

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Abstract

The Positivity Resonance Theory of co-experienced positive affect describes moments of interpersonal connection characterized by shared positive affect, caring nonverbal synchrony, and biological synchrony. The construct validity of positivity resonance and its longitudinal associations with health have not been tested. The current longitudinal study examined whether positivity resonance in conflict interactions between 154 married couples predicts health trajectories over 13 years and longevity over 30 years. We used couples' continuous ratings of affect during the interactions to capture co-experienced positive affect and continuous physiological responses to capture biological synchrony between spouses. Video recordings were behaviorally coded for co-expressed positive affect, synchronous nonverbal affiliation cues (SNAC), and behavioral indicators of positivity resonance (BIPR). To evaluate construct validity, we conducted confirmatory factor analysis to test a latent factor of positivity resonance encompassing co-experienced positive affect, co-expressed positive affect, physiological linkage of inter-beat heart intervals, SNAC, and BIPR. The model showed excellent fit. To evaluate associations with health and longevity, we used dyadic latent growth curve modeling and Cox proportional hazards modeling, respectively, and found that greater latent positivity resonance predicted less steep declines in health and increased longevity. Associations were robust when accounting for initial health symptoms, sociodemographic characteristics, health-related behaviors, and individually experienced positive affect. We repeated health and longevity analyses, replacing latent positivity resonance with BIPR, and found consistent results. Findings validate positivity resonance as a multimodal construct, support the utility of the BIPR measure, and provide initial evidence for the characterization of positivity resonance as a positive health behavior.

Keywords: broaden-and-build theory, marriage, positive psychology, affective science, dyadic interaction, health psychology

Positivity Resonance in Long-Term Married Couples:

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Although positive emotions often occur in connection with others – as spouses glance at each other lovingly, friends laugh together about an inside joke, or colleagues put their heads together to solve an intriguing research puzzle – the overwhelming majority of studies to date have examined positive emotions in individuals, using single-subject paradigms. Individuals who experience positive emotions reap many benefits (Fredrickson, 1998, 2001; Harker & Keltner, 2001; Isen, 2000; King et al., 2006; Tugade & Fredrickson, 2004). Prospective, longitudinal, and experimental intervention studies document that positive emotions contribute to well-being (Catalino & Fredrickson, 2011; Fredrickson et al., 2008; Lyubomirsky et al., 2005), health (Kok et al., 2013; Kok & Fredrickson, 2010; Pressman & Cohen, 2005; Richman et al., 2005), and even longevity (Chida & Steptoe, 2008; Diener & Chan, 2011). Affective scientists have only just begun to examine the unique moments of interpersonal connection that arise when one person’s positive emotional state simultaneously evokes – and is evoked by – another person’s positive emotional state. Grounded in *Positivity Resonance Theory* (Fredrickson, 2013, 2016), the present longitudinal study utilizes a rich dataset on long-term married couples. Our aim is to illuminate the characteristics and consequences of positivity resonance.

Positivity Resonance Theory of Co-Experienced Positive Affect

Drawing from both relationship and developmental science, Fredrickson (2016) proposed Positivity Resonance Theory as a generative way to study the emotion of love within affective science. In this framework, constructs commonly related to “love” (e.g., desire, intimacy, trust, commitments) are understood as products of the accumulation of momentary experiences of love, the emotion, defined as positivity resonance. Expanding on the broaden-and-build theory of

24 positive emotions (Fredrickson, 1998, 2001), moments of positivity resonance are taken to recur
25 between and among individuals and accumulate over time, functioning to build and fortify
26 enduring social bonds (love, the relationship) that later become steady resources for individuals
27 through good times and bad times (“in sickness and in health”). In other words, supportive social
28 bonds—together with their benefits for individuals’ health and well-being—emerge from a track
29 record of co-experienced positive affect (c.f., Gable et al., 2012).

30 *Theoretical Contributions*

31 Although emotions often occur in social contexts (e.g., Levenson, 2013; Smith et al.,
32 2004), most studies and theories in affective science focus on the emotions of one person. Even
33 in dyadic research, intraindividual affect often remains the unit of analysis (e.g., the extent to
34 which an individual's affect influences their partner's affect; Carstensen et al., 1995). Indeed, few
35 studies have focused on dyadic, linked emotional processes that transcend the individual (e.g.,
36 Levenson & Gottman, 1983; Timmons et al., 2015). Recently, theories of group-level affect have
37 emerged (Butler, 2017; Goldenberg et al., 2020), though they are rarely specific to group-level
38 *positive* affect. Positivity resonance addresses this theoretical gap in affective science by
39 highlighting the distinctive characteristics of co-experienced positive affect as well as its wide-
40 ranging contributions to health and well-being, including relationship health, public health,
41 and—our focus here—physical health and longevity (Brown & Fredrickson, 2021). Positivity
42 resonance itself may serve as a positive health behavior; yet no prior study has examined the
43 effects of positivity resonance on individual health and longevity.

44 Positivity Resonance Theory was inspired, in part, by prior work in relationship science
45 on perceived partner responsiveness (Reis, 2014), capitalization (i.e., sharing good news; Gable
46 & Reis, 2010) and expressed appreciation (Algoe et al., 2013). Positivity Resonance Theory

47 bridges affective science theory with relationship science theory by targeting holistic and
48 observable patterns of behavior emergent at the group level to offer a more general, cross-cutting
49 construct rooted in affective science. Complementing other seminal theories of relationship
50 science, Positivity Resonance Theory suggests an affective mechanism through which strong
51 attachments (Bowlby, 1969) and positive interdependence among individuals (Thibaut & Kelley,
52 1959) may occur. Positivity Resonance Theory calls for greater temporal precision to advance
53 scientific understanding of how momentary co-experiences of positive affect may ultimately
54 comprise the building blocks for broader relational constructs (e.g., trust, commitment,
55 relationship satisfaction).

56 *Characteristics*

57 Positivity resonance (Fredrickson, 2013, 2016) refers to moments of interpersonal
58 connection that arise when two or more individuals jointly experience positive emotions that are
59 elevated by the presence of key behavioral and physiological features. Consistent with how an
60 individual's experience of an emotion is coordinated across multiple response systems (i.e.,
61 experience, behavior, physiology; Levenson, 2014; Mauss et al., 2005; Wu et al., 2021),
62 moments of *positivity resonance* occur when two or more individuals engage in social interaction
63 characterized by three intertwined, collective responses: (a) shared positive affect (experiential),
64 (b) caring nonverbal synchrony (behavioral), and (c) biological synchrony (physiological).¹
65 Together, these three key features comprise the holistic experience of positivity resonance.

¹ Note that our current articulation of the three intertwined, defining features of positivity resonance has shifted slightly from its initial presentation (Fredrickson, 2013, 2016). Previously, the trio of collective responses was articulated as “(1) shared positive emotion, (2) mutual care, and (3) biobehavioral synchrony” (Fredrickson, 2016, p. 852). Our new phrasing decouples behavioral from biological synchrony to align better with the operationalized divisions among emotion response systems into experiential (i.e., shared positive-valence affect), behavioral (i.e., caring and synchronized nonverbal behaviors) and biological (i.e., physiological linkage) indicators, as has been done in recent articles (Brown & Fredrickson, 2021; Prinzing et al., 2020; West et al., 2021).

66 *Shared positive affect* refers to a pleasant subjective state that is jointly experienced
67 across multiple individuals. Although there are ways that positive affect can be potentially
68 maladaptive (e.g., too much, wrong context; Gruber et al., 2011), the biological, psychological,
69 and social benefits of positive affect are well-documented (e.g., Fredrickson et al., 2008; Kok et
70 al., 2013; Pressman & Cohen, 2005; Sin & Lyubomirsky, 2009). Positivity Resonance Theory
71 posits that these benefits are amplified when positive affect is shared between and among
72 individuals compared to when it is experienced individually. For example, in a laboratory study
73 in which romantic couples discussed how they first met, researchers coded the amount of time
74 spent laughing (either alone or simultaneously with their partner) from video recordings of the
75 conversations, and found that the proportion of time coded as shared laughter (independent of
76 time spent laughing alone) was associated with greater relationship quality, closeness, and social
77 support (Kurtz & Algoe, 2015). In large part, these additional benefits may emerge because
78 positive affect grows more intense and lasts longer when socially shared (e.g., Gable et al., 2004;
79 Kraut & Johnston, 1979). However, Positivity Resonance Theory suggests that even low
80 intensity shared positive affect yields more powerful benefits than does similarly intense positive
81 affect that is experienced individually (Fredrickson, 2016).

82 *Caring nonverbal synchrony* encompasses coordinated movements and gestures that
83 momentarily convey investment in the well-being of the other, a purported essential
84 characteristic of love (Hegi & Bergner, 2010). Momentary experiences of love, the emotion,
85 have been linked to four nonverbal affiliation cues: affirmative head nods, Duchenne smiles,
86 non-hostile hand gestures toward the other, and leaning toward the other, which signal approach
87 motivation, commitment, and trust (Gonzaga et al., 2001). Affiliation cues communicate care
88 and responsiveness to one's partner (Reis et al., 2004), which predict better relationship

89 outcomes (e.g., relationship well-being and longevity; Gable et al., 2006) and physical health
90 (e.g., lower mortality risk; Selcuk & Ong, 2013). Affiliation cues may also become mirrored and
91 synchronized into a “dance” of mutual attentiveness, positivity, and behavioral coordination
92 (Bernieri et al., 1988; Tickle-Degnen & Rosenthal, 1990; Vacharkulksemsuk & Fredrickson,
93 2012). A hallmark of positive interpersonal exchanges, behavioral synchrony can emerge as
94 early as infancy (i.e., between infants and their caregivers; Meltzoff & Moore, 1989) and can
95 occur cross-modally (i.e., beyond mimicry), such as when the rhythm of an infant’s movements
96 syncs up with the rhythm of a mother’s vocalizations (Stern et al., 1985). Laboratory studies of
97 adults show that synchronized body movements facilitate perceptions of embodied rapport
98 (Vacharkulksemsuk & Fredrickson, 2012), compassion (Valdesolo & Desteno, 2011), emotional
99 support satisfaction (Jones & Wirtz, 2007), and affiliation (Hove & Risen, 2009). Therefore, we
100 believe synchronized body movements that further indicate care, love, and affiliation (i.e., caring
101 nonverbal synchrony) represent a key component of high-quality moments of connection.

102 *Biological synchrony* occurs when biological response systems (e.g., physiological,
103 biochemical, neural) of two or more people change in coordinated ways. Consistent with
104 Positivity Resonance Theory, empirical evidence shows that biological synchrony emerges when
105 two or more people share a positive emotional state. For example, parent-infant pairs show
106 synchrony in oxytocin levels during mutual positive engagement (Feldman et al., 2010).
107 Neuroimaging studies also reveal widespread neural synchrony within dyads and groups sharing
108 a positive emotional experience (Hasson et al., 2004; Stephens et al., 2010). Synchrony in
109 autonomic physiology (also called “physiological linkage”) has been related to favorable
110 outcomes such as higher relationship quality (Helm et al., 2014), greater patient perceptions of
111 therapist empathy (Marci et al., 2007), and social bonding (for a review, see Feldman, 2015).

112 However, evidence for the association between physiological linkage and relationship outcomes
113 has been mixed (Timmons et al., 2015), which may reflect differing methods for measuring
114 linkage. For instance, early research on this topic found that greater *overall* physiological linkage
115 (a grand average measured over long time periods, e.g., across an entire 15-minute conversation)
116 was associated with adverse outcomes, such as lower marital satisfaction (Levenson & Gottman,
117 1983). For the present study, because we view emotions (e.g., love) as short-lived phenomena,
118 we consider *momentary* physiological linkage during seconds characterized by shared positive
119 affect to be a more appropriate method for capturing biological synchrony, rather than overall, or
120 *grand average*, linkage (Chen et al., 2020; described more fully below).

121 *Associations Among Defining Features*

122 Each of the three defining features of positivity resonance is theoretically aligned with a
123 particular emotion response system (i.e., subjective experience, behavior, physiology).
124 According to a number of emotion theorists, emotions involve coordinated changes across these
125 response systems, a process often referred to as *emotion coherence* (Ekman, 1992; Levenson,
126 1994). Building on this idea of within-person emotion coherence, Positivity Resonance Theory
127 suggests that high-quality moments of connection involve further coordination, occurring *across*
128 individuals, as reflected by the co-occurrence of its three key features. Although research has
129 evaluated emotional responding across individuals in the same response system (e.g., emotional
130 convergence of subjective experience, synchrony in physiological responses; Anderson et al.,
131 2003; Levenson & Gottman, 1983), less is known about the multimodal, interpersonal emotion
132 coherence that is theorized to occur during moments of positivity resonance.

133 *Consequences*

134 Preliminary evidence suggests that positivity resonance may promote health and well-
135 being. Initial research on the consequences of positivity resonance found that participants who,
136 over a 9-week period of nightly self-reports, showed increases in feeling “close” and “in tune”
137 with others (a possible proxy for shared positive affect and caring nonverbal synchrony) had
138 increases in cardiac vagal tone (Kok & Fredrickson, 2010), which is correlated with physical
139 health (Bibeovski & Dunlap, 2011; Thayer & Sternberg, 2006). The first published empirical
140 research on positivity resonance validated a new self-report measure of perceived positivity
141 resonance and showed that it is associated, within individuals, with flourishing mental health,
142 fewer depressive symptoms, loneliness, and (albeit less consistently) illness symptoms, even
143 when controlling for daily pleasant emotions or amount of social interaction more generally
144 (Major et al., 2018). More recent research that used this same measure of perceived positivity
145 resonance during the early months of the COVID-19 pandemic found it to account for the link
146 between trait resilience and mental health (Prinzing et al., 2020) and also to predict behaviors
147 known to promote public health (i.e., handwashing, mask wearing, and social distancing), as
148 mediated by prosocial tendencies (West et al., 2021). Here, we aim to advance this prior work by
149 measuring positivity resonance through a suite of objective, dyad-level methods and in a social
150 context (i.e., long-term marriage) to further illuminate its longitudinal consequences for health
151 and longevity.

152 **Development of Objective and Dyad-Level Measures of Positivity Resonance**

153 *Longitudinal Study of Long-Term Married Couples*

154 Through a series of studies using data from the same dataset analyzed here, we developed
155 and validated new, dyad-level measures of positivity resonance using multiple methods. This
156 dataset draws from an unparalleled longitudinal study of middle-aged and older couples in long-

157 term marriages (Levenson et al., 1993; Levenson et al., 1994). In the first laboratory session of
158 the study, couples engaged in three 15-min conversational interactions: (a) a discussion of the
159 events of the day, (b) a discussion of an area of continuing disagreement in their marriage, and
160 (c) a discussion of a mutually agreed upon pleasant topic. For the present study, we measured
161 positivity resonance during the discussion of a disagreement (i.e., conflict conversation), a
162 context that is familiar to most couples and one that is rich with not only negative but also
163 positive emotion (Haase et al., 2013; McGonagle et al., 1992), to maximize ecological validity.

164 *Defining Features*

165 In our study of *shared positive affect*, we utilized each spouse's moment-by-moment
166 ratings of their individual affective experience during the interactions, which they provided by
167 continuously moving a rating dial while watching the video-recordings of their interactions. We
168 found that co-experienced positive affect (the number of seconds in which both spouses reported
169 feeling positive), more than individually experienced positive affect (the number of seconds in
170 which one partner reported feeling positive and the other did not) was associated with greater
171 marital satisfaction (Brown et al., 2021). In another study, we measured *caring nonverbal*
172 *synchrony* during the conflict interaction by applying a dyad-level modification to a behavioral
173 coding system developed by Gonzaga and colleagues (2001), coding the same nonverbal
174 affiliation cues (e.g., head nods, smiles) that have been associated with love (versus desire), yet
175 with exclusive focus on those occurring synchronously (i.e., both partners displayed an
176 affiliation cue near simultaneously). Preliminary analyses suggest that synchronized nonverbal
177 affiliation cues are positively associated with wives' perceptions of husbands' lovingness (Lai et
178 al., in prep).

179 In our study of *biological synchrony*, we measured physiological linkage over short time
180 periods (i.e., 15-second rolling time windows) in the conflict interaction during four emotion
181 categories defined by behavioral coding: co-expressed positive emotion, co-expressed negative
182 emotion, co-expressed neutral emotion (i.e., both showed no emotion), and individually
183 expressed emotion (Chen et al., 2020). Results revealed that co-expressed positive emotion,
184 relative to all other emotion categories, is associated with greater in-phase physiological linkage
185 (responses changing in the same direction) and lower anti-phase physiological linkage (responses
186 changing in opposite directions). Greater in-phase physiological linkage during co-expressed
187 positive emotion was also positively associated with the overall affective quality of the
188 interaction and marital satisfaction (Chen et al., 2020). Further, the *momentary* physiological
189 linkage approach outperformed the *grand average* approach (i.e., measuring linkage across the
190 entire conversation) in its associations with related constructs like affective and marital quality,
191 and thus appears to be a more useful measure for evaluating positivity resonance.

192 ***Holistic Measure***

193 Positivity Resonance Theory suggests that its three defining features may combine
194 synergistically and be particularly powerful when they co-occur, rather than when they occur
195 separately. Motivated by this hypothesis, we created a novel, group-level measure of behavioral
196 indicators of positivity resonance (BIPR) that integrates multiple features of positivity resonance
197 (e.g., shared positive affect, mutual care and concern, and behavioral synchrony). This
198 behavioral coding system combines actions, words, and voice intonation that convey mutual
199 warmth, concern, affection and/or a shared tempo into one holistic measurement of positivity
200 resonance. In the initial study of BIPR, we found that it is a more potent predictor of marital

201 satisfaction than a behavioral measure of co-expressed positive affect alone (i.e., without
202 consideration of mutual care or behavioral synchrony; Otero et al., 2019).

203 In sum, we have begun to examine positivity resonance, its characteristics, and correlates
204 using the present longitudinal dataset of long-term married couples. Importantly, no prior study
205 has investigated the multimodal construct validity of positivity resonance nor its longitudinal
206 associations with health and longevity. Moreover, important unanswered questions remain
207 regarding for whom (e.g., women versus men) positivity resonance may be the most beneficial
208 and how it is best assessed (e.g., using one or multiple measures).

209 **Additional Questions**

210 *Gender Differences*

211 Positivity resonance is a group-level phenomenon (Fredrickson, 2016), and is thought to
212 be beneficial to all those who experience it. However, gender-specific effects are common in
213 marital research on heterosexual couples (Baucom et al., 1990; Kiecolt-Glaser & Newton, 2001).
214 Evidence is mixed regarding whether the effects of relationships on health are stronger for
215 women versus men, including studies using the same longitudinal dataset as used here (Bloch et
216 al., 2014; Haase et al., 2016; Levenson et al., 1993), as well from other studies. For example, a
217 15-year study using medical records found relationship characteristics (e.g., companionship,
218 equality in decision-making) to be associated with a lower risk of death in married women, but
219 not men (Hibbard & Pope, 1993). At the same time, evidence supports the opposite conclusion,
220 that men's health may be more closely tied to aspects of the marriage. Laboratory studies of
221 marital conflict have linked hostility with heightened cardiovascular reactivity (Smith & Gallo,
222 1999); anger with increases in blood pressure (Miller et al., 1999); and stonewalling with lower
223 physical health (Gottman, 1991) – for men, in particular.

224 A large body of evidence also points to gender differences in emotion and social
225 relationships. Women tend to be more emotionally expressive than men (for a review, see Brody
226 & Hall, 2000), as measured by observational coding (e.g., Kring & Gordon, 1998) and facial
227 electromyography (e.g., Bradley et al., 2001). Compared to men, women smile more when
228 engaging with others (LaFrance et al., 2003) and express more voiced laughter (Bachorowski et
229 al., 2001), which elicits more positive affect in listeners than unvoiced laughter (Bachorowski &
230 Owren, 2001). Additionally, women have larger social networks compared to men and are more
231 likely to maintain active friendships throughout their lives (Candy et al., 1981; Field & Minkler,
232 1988). These patterns suggest that women may have more opportunities for social interactions
233 than men—and may be more likely than men to express positive affect, experience positive
234 affect themselves, and to elicit positive affect in their interaction partners. Conceivably, women
235 who tend to cultivate positivity resonance in their marriage may also do so in other social
236 relationships, potentially resulting in higher overall “doses” of positivity resonance for wives,
237 compared to their husbands.

238 Given somewhat inconsistent evidence for gender differences in the scientific literatures
239 on marriage, emotion, and relationships, we did not have a specific hypothesis regarding whether
240 couples’ positivity resonance may be more important for wives’ or husbands’ health and
241 longevity. Positivity Resonance Theory also makes no predictions about gender differences.
242 Thus, we explored this question in the present study.

243 *Measurement Parsimony*

244 Positivity Resonance Theory proposes that the combination of shared positive affect,
245 caring nonverbal synchrony, and biological synchrony promote long-term health outcomes,
246 above and beyond any single feature in isolation. However, given the practical constraints of

247 many research settings, it may not be possible to assess all these features simultaneously. Thus,
248 the holistic behavioral measure of positivity resonance, BIPR, may be a useful tool for
249 researchers with more limited resources. It remains to be determined whether BIPR would
250 perform as well as a comprehensive latent factor that incorporates multiple measures and
251 features of positivity resonance in predicting long-term health and longevity.

252 **Long-Term Marriage as a Context for Studying Positivity Resonance**

253 Marriages are among the most significant relationships in adult life (more than 94% of
254 U.S. Americans over the age of 55 have been married at least once; U.S. Bureau of the Census,
255 2011). Marriages may be especially significant in later life as social networks shrink and close
256 relationships become increasingly important (Carstensen et al., 1999). A long line of research has
257 evaluated characteristics of marriages that are associated with different health-related outcomes,
258 with particular focus on spouses' emotional functioning (Gottman & Levenson, 1986; Levenson
259 et al., 2013; Smith et al., 2014), which is known to have downstream consequences for well-
260 being (e.g., Carr et al., 2014; Glenn & Weaver, 1981), mental health (e.g., Beach, 2014; Beach et
261 al., 1998), and physical health (e.g., Haase et al., 2016; Kiecolt-Giaser et al., 1993; Robles &
262 Kiecolt-Giaser, 2003). Much of this work was devoted to uncovering negative emotional
263 qualities of marriages and their consequences (e.g., Gottman & Levenson, 1992; Kiecolt-Giaser
264 et al., 1993). More recently, another line of research has emerged documenting the positive
265 emotional qualities of marriage and close relationships (e.g., Algoe et al., 2013; Gable et al.,
266 2004; Laurenceau et al., 2005), and the consequences these positive qualities have, independent
267 of the adverse effects of negative emotions (e.g., Algoe, 2019; Feeney & Collins, 2015;
268 Pietromonaco & Collins, 2017).

269 Emotions may have especially long-lasting consequences, such as predicting longevity, in
270 the context of long-term marriage, given the duration and importance of this relationship. Indeed,
271 individuals who rate their marriage as happier have significantly lower odds of dying (Whisman
272 et al., 2018). Beyond intrapersonal associations among emotion, marriage, and longevity, there is
273 also evidence that having a happier spouse predicts greater longevity in elderly couples
274 (Stavrova, 2019). Moreover, greater self-reported perceived partner responsiveness (i.e., a key
275 feature of caring nonverbal synchrony) has been linked with lower all-cause mortality in
276 romantic couples (Selcuk & Ong, 2013; Stanton et al., 2019). Additional research is needed to
277 explore whether interpersonal emotional processes at the level of the dyad (e.g., positivity
278 resonance) predict health and longevity in long-term married couples; and further, whether these
279 predictions are independent of individual-level emotions, marital quality, or both.

280 **Importance of Longitudinal Assessment**

281 It is important to utilize a longitudinal design when studying associations between
282 emotions and health. Positivity resonance may well be linked to present-day health and well-
283 being, as is suggested (albeit inconsistently) by Major et al. (2018). Yet, we expect its effects
284 may be amplified throughout the course of relationships, as moments of positivity resonance
285 recur and accumulate over time (Fredrickson, 1998, 2001). The effects of relationships on health
286 may also become stronger over time, as individuals age and their social networks shrink (Rook &
287 Charles, 2017). Additionally, health is known to decline with age (Pinquart, 2001), and positivity
288 resonance may promote health longitudinally by protecting against these normative declines in
289 health. Therefore, health consequences of positivity resonance in marriages may be more evident
290 longitudinally than cross-sectionally. For examining these kinds of questions, longitudinal
291 designs clearly have advantages over more common cross-sectional designs.

292 **The Present Study**

293 Using the present rich, longitudinal dataset (Levenson et al., 1993; Levenson et al.,
294 1994), we have recently developed novel, objective, dyad-level measures of positivity resonance
295 (i.e., Brown et al., 2021; Chen et al., 2020; Lai et al., in prep; Otero et al., 2019). The present
296 study had two aims: (1) to examine the covariance among these measures through a
297 measurement model of positivity resonance as a single latent factor (i.e., through confirmatory
298 factor analysis [CFA]); and (2) to use this latent factor to predict longitudinal health trajectories
299 and longevity.

300 To pursue our first aim, we conducted CFA to test a measurement model of positivity
301 resonance, indicated by our dyad-level measures. Our first hypothesis (Hypothesis 1) was that
302 the CFA would fit satisfactorily, supporting the existence of a broader positivity resonance
303 construct with multimodal manifestations of its defining features, objectively assessed at the
304 dyadic level, in the domains of experience, behavior, and physiology. To pursue the second aim,
305 we conducted two series of analyses examining whether couples' positivity resonance (measured
306 at the first timepoint) predicted (a) longitudinal trajectories of wives' and husbands' health
307 symptoms (measured at three timepoints, separated by approximately 6-7 years) as well as (b)
308 mortality (measured over the ensuing 20 years). We hypothesized that greater positivity
309 resonance would be associated with less steep declines in health (Hypothesis 2) and increased
310 longevity (Hypothesis 3) in both wives and husbands. Analyses for Hypotheses 2 and 3
311 proceeded in five steps: (1) We conducted preliminary analyses to verify selection of model
312 parameters; (2) We examined associations of our latent factor of positivity resonance with health
313 trajectories (controlling for health at T1) and longevity, respectively; (3) We explored whether
314 gender moderated associations of latent positivity resonance with health and longevity; (4) We

315 examined the robustness of our findings by controlling for (a) sociodemographic factors (e.g.,
316 age, education), behaviors known to influence health (e.g., smoking, exercise), and individually
317 experienced positive affect during the conflict discussion at the first timepoint (to investigate the
318 added value of dyad-level, co-experienced positive affect, independent of individually
319 experienced positive affect), and (b) marital satisfaction; and (5) Finally, to examine whether
320 associations with health trajectories and longevity could be obtained with a single behavioral
321 measure of positivity resonance, we repeated longitudinal analyses replacing our latent factor of
322 positivity resonance, as the independent variable, with BIPR.

323 **Method**

324 **Participants**

325 We analyzed archival data from a longitudinal study of 156 heterosexual long-term
326 married couples. The current sample ($N = 154$ couples; $n = 2$ couples excluded due to missing
327 data) was comprised of a middle-aged cohort ($n = 80$ couples; M age = 44.33 years; SD age =
328 2.92 years) and an older adult cohort ($n = 74$ couples; M age = 63.54 years; SD age = 3.21 years).
329 The sample was recruited from the San Francisco Bay Area to be representative of the
330 demographic characteristics (socioeconomic status, religion, ethnicity) of couples in these age
331 groups in that area at the time of the study. The resulting sample was primarily white (86%),
332 Protestant or Catholic (62%), relatively well-off socioeconomically, and with children (96% of
333 couples had at least one child). Complete details of the sampling and recruitment procedures
334 have been reported previously (Levenson et al., 1993). Several prior studies have analyzed data
335 from this sample (e.g., Bloch et al., 2014; Brown et al., 2021; Chen et al., 2020; Haase et al.,
336 2016; Otero et al., 2019; see **Previous Publications** and **Supplemental References**, Online
337 Supplemental Materials), mostly focusing on the early waves of assessment. However, no prior

338 studies have examined longitudinal associations between positivity resonance and health or
339 longevity.

340 **Procedure**

341 Data were initially collected at three time points over the course of approximately 13
342 years (Time 1 (T1): 1989/90, $N = 154$ couples; Time 2 (T2): 1995/96, $n = 131$ couples; Time 3
343 (T3): 2001/02, $n = 101$ couples). Longevity data were collected during a follow-up phase 30
344 years later, spanning from June 1, 2020 to April 1, 2021. Attrition in the sample occurred when
345 couples discontinued participation for the following reasons (cumulative frequencies): (a)
346 divorce (T2: $n = 5$; T3: $n = 8$); (b) death of a spouse (T2: $n = 10$; T3: $n = 25$), or (c)
347 declined/unknown reasons (T2: $n = 8$; T3: $n = 21$). We also examined whether health symptoms
348 and positivity resonance were associated with drop-out. Health symptoms at T1 did not predict
349 drop-out over time. Positivity resonance at T1 was associated with drop-out at T3, $t(147.54) =$
350 $5.36, p < .001$); couples who discontinued the study at T3 had lower positivity resonance ($M = -$
351 $0.46, SD = 0.63$) than those who continued in the study ($M = 0.24, SD = 1.01$). We used full
352 information maximum likelihood estimation (FIML; e.g., Jeličić et al., 2009) to account for
353 missing data in the CFA and throughout the longitudinal health trajectory analyses.

354 At each time point, couples completed questionnaires and participated in a laboratory
355 session that followed well-established procedures for studying marital interactions (Levenson &
356 Gottman, 1983). Couples engaged in three 15-minute conversations: (a) events of the day (T1) or
357 events since the last assessment (T2 and T3); (b) conflict topic – an issue of ongoing
358 disagreement in their marriage; and (c) pleasant topic – something they enjoyed doing together.
359 The present study analyzed data from the conflict conversation only.

360 Partially hidden cameras were used to videotape each interaction for subsequent

361 behavioral coding (see below). Several days after each laboratory session, each participant
362 returned to the laboratory to watch video-recordings of their conversations, individually, while
363 providing continuous ratings of how they felt during the interactions using a rating dial,
364 consisting of small metal box with a rotating pointer that traversed a 180° path (a well-validated
365 procedure for obtaining continuous self-reported affect; Gottman & Levenson, 1985).
366 Participants continuously moved the rating dial across a 9-point scale anchored by the legends
367 “extremely negative” (1) and “extremely positive” (9), with a line labeled “neutral” in the middle
368 (5). The dial generated a voltage that reflected the dial position; a computer sampled the voltage
369 100 times per second, and computer software developed by Robert W. Levenson computed the
370 average dial position every second.

371 Couples’ physiological responses were recorded continuously throughout all interactions
372 using a Grass Model 7 12-channel polygraph and the same computer that was used for sampling
373 rating dial voltage (described above). For the present study, we focus on linkage in inter-beat
374 intervals (IBI) of the heart, because this physiological channel showed the highest effect sizes in
375 the original study of physiological linkage (Chen et al., 2020), relative to the other physiological
376 indices, and as such appears to be more sensitive to changes in dyadic emotion. Cardiac IBI was
377 obtained using Beckman miniature electrodes with Redux paste that were placed in a bipolar
378 configuration on opposite sides of the participant's chest. IBI was measured as the interval
379 between successive R-waves of the electrocardiogram was measured in milliseconds.

380 All procedures were approved by the University of California, Berkeley Committee for
381 the Protection of Human Subjects. This study was not preregistered.

382 **Measures**

383 *Positivity resonance (TI)*

384 Couples' positivity resonance was modeled as a latent variable, indicated by five dyad-
385 level measures (each measure is listed as a subheading and described below). Each dyad-level
386 measure was calculated across the entire 15-minute conflict conversation to obtain one value for
387 each couple, such that all measures are temporally comparable and reflect the same time period.
388 Descriptive statistics, sample sizes, and intercorrelations among dyad-level positivity resonance
389 variables are provided in **Table 1**.

390 **Table 1**

391 *Descriptive Statistics and Intercorrelations Among Positivity Resonance Latent Factor and its Dyad-Level Indicators*

Variables	1	2	3	4	5	Mean	SD	Min	Max	<i>n</i>
1. Positivity resonance (factor scores)						0	0.96	-1.09	3.53	154
2. BIPR	0.98***					5.92	5.89	0	32	148
3. SNAC	0.87***	0.79***				12.3	10.37	0	46	147
4. Co-expressed positive affect	0.64***	0.56***	0.52***			26.44	31.01	0	149	150
5. Co-experienced positive affect	0.28**	0.25**	0.16*	0.29***		260.91	229.6	0	900	153
6. In-phase IBI linkage ^a	0.26**	0.26**	0.19*	0.08	-0.01	0.44	0.23	0	0.99	114

Note. BIPR = Behavioral Indicators of Positivity Resonance. SNAC = Synchronized Nonverbal Affiliation Cues. SD = standard deviation. Min = minimum. Max = maximum. IBI = inter-beat interval. ^aduring co-expressed positive affect. * $p < .05$. ** $p < .01$. *** $p < .001$.

392

393 **Behavioral Indicators of Positivity Resonance (BIPR).** Couples' behavior was coded
394 using a dyad-level coding system (Otero et al., 2019) that captured holistic, integrated behavioral
395 indications of positivity resonance using the following prompt: "*Did positivity resonate between*
396 *the two partners? That is, did they show actions, words, or voice intonation that conveyed*
397 *mutual warmth, mutual concern, mutual affection and/or a shared tempo (i.e., shared smiles and*
398 *laughter)?*" Three trained coders viewed the videotaped conflict interactions and rated BIPR
399 every 30 seconds on a 3-point intensity scale (0 = not present; 1 = lower intensity or present
400 once; and 2 = higher intensity or present more than once). Coders did not evaluate the presence
401 of negative emotional behaviors in their BIPR ratings; that is, negative emotional behaviors were
402 not weighted against indications of positivity resonance that occurred in the same coding period.
403 To assess interrater reliability, all three coders coded 20% of the study sample. Reliability was
404 high (intraclass correlation coefficient = .80). Codes were summed across all 30-second periods
405 to obtain one BIPR score for the entire conversation.

406 **Synchronized Nonverbal Affiliation Cues (SNAC).** Couples' synchronized nonverbal
407 affiliation cues (i.e., caring nonverbal synchrony) were assessed using a recently developed
408 behavioral coding system that captures simultaneous or near-simultaneous nonverbal affiliation
409 cues between partners (Lai et al., in prep). SNAC is based on a coding system that captures four
410 nonverbal displays of love/affiliation at the individual-level (e.g., head nods, smiles, forward
411 leans, and non-hostile hand gestures; Gonzaga et al., 2001). An independent team of trained
412 coders (i.e., different coders than those who coded BIPR) viewed the videotaped conflict
413 interactions, without audio, and rated SNAC every 30 seconds on a 0-2 scale. Again, coders did
414 not take into consideration expressions of negative emotional behaviors (e.g., frowns). Codes

415 were averaged across coders (reliability was high; intraclass correlation coefficient = .86-.90)
416 and summed across all 30-second periods to obtain one SNAC score for the entire conversation.

417 **Co-Expressed Positive Affect.** Each spouse's emotional behavior was coded using the
418 Specific Affect Coding System (SPAFF; Coan & Gottman, 2007), which evaluated verbal
419 content, voice tone, context, facial expression, gestures, and body movements. There are five
420 positive speaker codes (interest, affection, humor, validation, joy), nine negative speaker codes
421 (anger, contempt, disgust, belligerence, domineering, defensiveness, fear/tension/worry, sadness,
422 whining), and three listener emotion codes (positive, negative, stonewalling). An independent
423 team of trained coders viewed the videotaped conflict interactions and rated each spouse's
424 emotional behaviors on a second-by-second basis. For both speakers and listeners, a "neutral"
425 code (0 = absent, 1 = present) was assigned for seconds in which neither positive nor negative
426 emotional behavior were coded. Interrater reliability of the SPAFF coding was satisfactory
427 (overall mean kappa = .64). Additional details regarding SPAFF reliability in this sample has
428 been published elsewhere (Carstensen et al., 1995). Co-expressed positive affect was calculated
429 for each couple as the number (sum) of seconds in which both partners were simultaneously
430 coded with a positive SPAFF code (i.e., either as a speaker or listener; regardless of intensity). In
431 other words, this measure is specific to the *cumulative duration* of co-expressed positive affect
432 and does not take intensity into consideration. In addition, moments of individually experienced
433 positive affect (i.e., seconds in which one partner expresses positive affect while the other
434 partner expresses negative or neutral affect) are not counted towards this variable.

435 **Co-Experienced Positive Affect.** The average rating dial position for each spouse's
436 ratings of how they felt during the conflict interaction was computed for every second.
437 Following data reduction procedures from the validation study of shared positive affect, couples'

438 co-experienced positive affect was recorded as the number (sum) of seconds in which both
439 partners reported experiencing positive affect (≥ 5 or “neutral”² on the rating dial at the same
440 time; Brown et al., 2021). Again, this measure is specific to the *cumulative duration* of co-
441 experienced positive affect, regardless of intensity, and only includes co-experienced, rather than
442 individually experienced, positive affect.

443 **In-Phase IBI Linkage.** IBI data for the conflict interaction were averaged every second
444 and smoothed using a 10-second rolling time window. For each couple, a time series of total IBI
445 linkage was computed by calculating Pearson’s correlations between both partners’ IBI
446 responses within 15-second rolling time windows (Marci et al., 2007; Marci & Orr, 2006). We
447 then computed a time series of in-phase IBI linkage by entering the correlation coefficient from
448 the total linkage time series if it was positive or entering a 0 if the correlation was 0 or negative.
449 In the present study, we examine momentary in-phase IBI linkage during moments of co-
450 expressed positive affect, given its previous association with marital satisfaction (Chen et al.,
451 2020). We calculated the average degree of in-phase IBI linkage during moments of co-
452 expressed positive affect by taking the average level of in-phase IBI linkage across all seconds
453 where both partners were simultaneously coded with a positive SPAFF code (see above).

454 ***Health Symptoms (T1, T2, T3)***

455 Health symptoms were measured using the Cornell Medical Index (CMI; Brodman et al.,
456 1949). The CMI is a well-established self-report measure that contains 195 items assessing a
457 variety of mental and physical health symptoms. The CMI shows high convergence with medical

² As in the validation study (Brown et al., 2021), we included the neutral line (5 on the rating dial) in the threshold for determining positive affect because (a) Positivity Resonance Theory posits that even low intensity co-experienced positive affect is beneficial (Fredrickson, 2016) and (b) given the nature of the rating dial, participants necessarily move through the neutral line in order to shift from negative to positive affect, without necessarily feeling neutral.

458 evaluations of health and predicts morbidity over time (Weaver et al., 1980). Because we wanted
459 to focus on current health, we excluded 13 CMI items that assessed family history of illness and
460 5 items assessing behaviors known to influence health, such as smoking and drinking (as has
461 been done in previous studies using the CMI; e.g., Aldwin et al., 1989; Aldwin et al., 2001;
462 Haase et al., 2016). To reduce skew, items were recoded (0 = symptom not present; 1 = symptom
463 present [regardless of intensity]) following established procedures (e.g., Duncan et al., 2006;
464 Haase et al., 2016). A total health symptoms score was calculated at each timepoint by taking the
465 sum of all items (excluding family history and health-related behaviors). Lower scores on the
466 CMI indicate better health, with 0 representing no symptoms and 177 representing the highest
467 total possible score. Descriptive statistics for health symptoms and covariates are presented in
468 **Table 2.**

469 **Table 2**

470 *Descriptive Statistics for Key Individual-Level Study Variables*

	Wives		Husbands	
	Mean (SD)	Range	Mean (SD)	Range
T1 health symptoms	18.82 (14.95)	2 - 129	13.47 (8.53)	0 - 50
T2 health symptoms	19.76 (13.08)	2 - 82	14.23 (9.60)	0 - 56
T3 health symptoms	18.76 (11.58)	2 - 56	14.67 (9.22)	0 - 61
Age	52.91 (10.03)	37 - 70	54.21 (10.17)	39 - 70
Household income (<i>n</i> =)				
less than \$10,000		1		1
\$10,000 - \$19,999		3		3
\$20,000 - \$29,999		6		6
\$30,000 - \$39,999		16		16
\$40,000 - \$49,999		25		25
\$50,000 - \$59,999		28		28
\$60,000 - \$69,999		23		23
\$70,000 - \$79,999		14		14
\$80,000 - \$89,999		14		14
\$90,000 - \$99,999		6		6
\$100,000 or more		17		17
Education	23.44 (7.10)	8 - 34	26.46 (7.38)	10 - 35
Health-related behaviors	0.80 (0.90)	0 - 4	0.74 (0.88)	0 - 3
Individual ^a PA	208.83 (208.58)	0 - 900	170.47 (178.69)	0 - 869
Marital satisfaction	111.3 (16.91)	46.5 - 138	111.3 (17.08)	43.5 - 138

Note. Household income is a dyad-level covariate; values are the same across wives and husbands. ^aIndividually experienced. PA = positive affect. SD = Standard deviation; T1 = Time 1; T2 = Time 2; T3 = Time 3.

471

472 ***Mortality***

473 Between the beginning of the study in 1989 and the start of the search period for
474 collecting mortality data (June 1, 2020), 135 deaths were confirmed (43.8%). Deceased

475 participants' date of death was obtained from report of relatives ($n = 36$), the United States
476 Social Security Death Index database (<http://ssdi.genealogy.rootsweb.com>; $n = 54$); online
477 obituary listings ($n = 36$), or through another online search engine (i.e., facebook.com,
478 intelius.com; $n = 9$), following procedures used successfully in previous studies collecting
479 longitudinal or mortality data (e.g., Bolanos et al., 2012; Engoren et al., 2002; Shelton et al.,
480 2018; van Kimmenade et al., 2010). Survival time was computed as the number of days between
481 the date of their initial laboratory visit and the date of death. We confirmed that 145 participants
482 (45.1%) were still alive after June 1, 2020 through phone/email contact with participants and
483 their relatives ($n = 136$) and social media (e.g., facebook.com, linkedin.com; $n = 9$). Data from
484 participants who had not died (i.e., their exact survival time is unknown) were censored, a
485 common data estimation technique used in survival analysis when the event of interest has not
486 yet occurred (Finkelstein, 1986). Censor time for these participants was computed as the number
487 of days between the date of their initial laboratory visit and June 1, 2020 (Leon et al., 1990). For
488 the remaining 28 participants (9.5%) whose status was not confirmed within our search period
489 (06/01/2020 and 04/01/2021), censor time was conservatively computed as the number of days
490 between the date of their initial laboratory visit and their last known date alive (i.e., last
491 laboratory visit or questionnaire completion).³

492 *Covariates (T1)*

493 **Sociodemographic Characteristics.** Sociodemographic characteristics included age (in
494 years), annual household income before taxes (coded: 0 = less than \$10,000; 1 = \$10,000 -
495 \$19,999; 2 = \$20,000 - \$29,999; 3 = \$30,000 - \$39,999; 4 = \$40,000 - \$49,999; 5 = \$50,000 -

³ When we exclude participants whose living status is unknown ($n = 28$) from mortality analyses, results were consistent.

496 \$59,999; 6 = \$60,000 - \$69,999; 7 = \$70,000 - \$79,999; 8 = \$80,000 - \$89,999; 9 = \$90,000 -
497 \$99,999; and 10 = \$100,000 or more), and education (in years).

498 **Health-Related Behaviors.** Health-related behaviors included smoking (≥ 20 cigarettes
499 per day), alcohol consumption (≥ 2 drinks a day), caffeine consumption (≥ 6 cups of coffee or tea
500 per day), and lack of physical exercise from the CMI (recoded as 0 = no, 1 = yes) and summed.

501 **Individually Experienced Positive Affect.** Individually experienced positive affect was
502 determined separately, for wives and husbands, as the number (sum) of seconds in which the
503 individual reported experiencing positive affect (≥ 5 on the rating dial), while their partner did
504 not.

505 **Marital Satisfaction.** Marital satisfaction was assessed using two well-validated self-
506 report inventories: (a) the 15-item Marital Adjustment Test (Locke & Wallace, 1959), which
507 assesses agreement between spouses in various life domains (e.g., handling family finances,
508 demonstrations of affection); and (b) the 22-item Marital Relationship Inventory (Burgess et al.,
509 1971), which measures satisfaction with affection and sexuality in the marriage, overall
510 satisfaction with the marriage, and areas of agreement (e.g., “How happy would you rate your
511 marriage?”). Consistent with previous research (e.g., Carstensen et al., 1995) and to reduce Type
512 I error, we averaged the measures separately for husbands and wives to capture each spouse’s
513 marital satisfaction.

514 **Analytic Approach**

515 The present study used subjective experiential, behavioral, and physiological data
516 obtained during the conflict conversation at T1 to measure positivity resonance; self-reported
517 questionnaire data obtained at T1, T2, and T3; and mortality data obtained between June 1, 2020
518 and April 1, 2021 (see above). Preliminary CFA and longitudinal health trajectory analyses were

519 conducted within a structural equation modeling (SEM) framework, employing FIML to handle
520 missing data, through the *lavaan* package in R Studio Version 1.2.1335 (Rosseel, 2012). To
521 evaluate model fit in SEM, we inspected the χ^2 test of model fit as an absolute fit index as well as
522 the comparative fit index (CFI) and standardized root mean squared residual (SRMR) as relative
523 fit indices, following established guidelines (Hu & Bentler, 1999). Nonsignificant χ^2 values ($ps >$
524 0.05); CFI values greater than 0.95 and SRMR values less than .08 were used to indicate
525 satisfactory model fit. Mortality analyses were conducted using the *survival* package (v3.2-11;
526 Therneau, 2020). All continuous variables were standardized before analysis.

527 *Preliminary Analyses*

528 First, we examined intercorrelations among dyad-level variables (see **Table 1**) and
529 individual-level variables (see **Table 2**). Next, we conducted analyses to validate the assessment
530 of our key constructs (i.e., positivity resonance, health trajectories).

531 **Positivity resonance.** We evaluated the construct validity of positivity resonance, a
532 dyad-level latent variable indicated by an a priori set of observed indicator variables, using CFA.
533 We tested a measurement model of positivity resonance based on the following dyad-level
534 indicator variables: BIPR, SNAC, co-expressed positive affect, co-experienced positive affect,
535 and average in-phase IBI linkage during co-expressed positive affect. To reduce the number of
536 parameters, we factor scored the latent positivity resonance variable to obtain model-implied
537 values (i.e., weighting observed values based on parameter estimates and standardizing) for use
538 in all subsequent analyses (DiStefano et al., 2009).

539 **Health Trajectories.** We constructed a series of latent growth curve models (LGMs;
540 Olsen & Kenny, 2006) with latent intercepts and slopes of health trajectories for husbands and

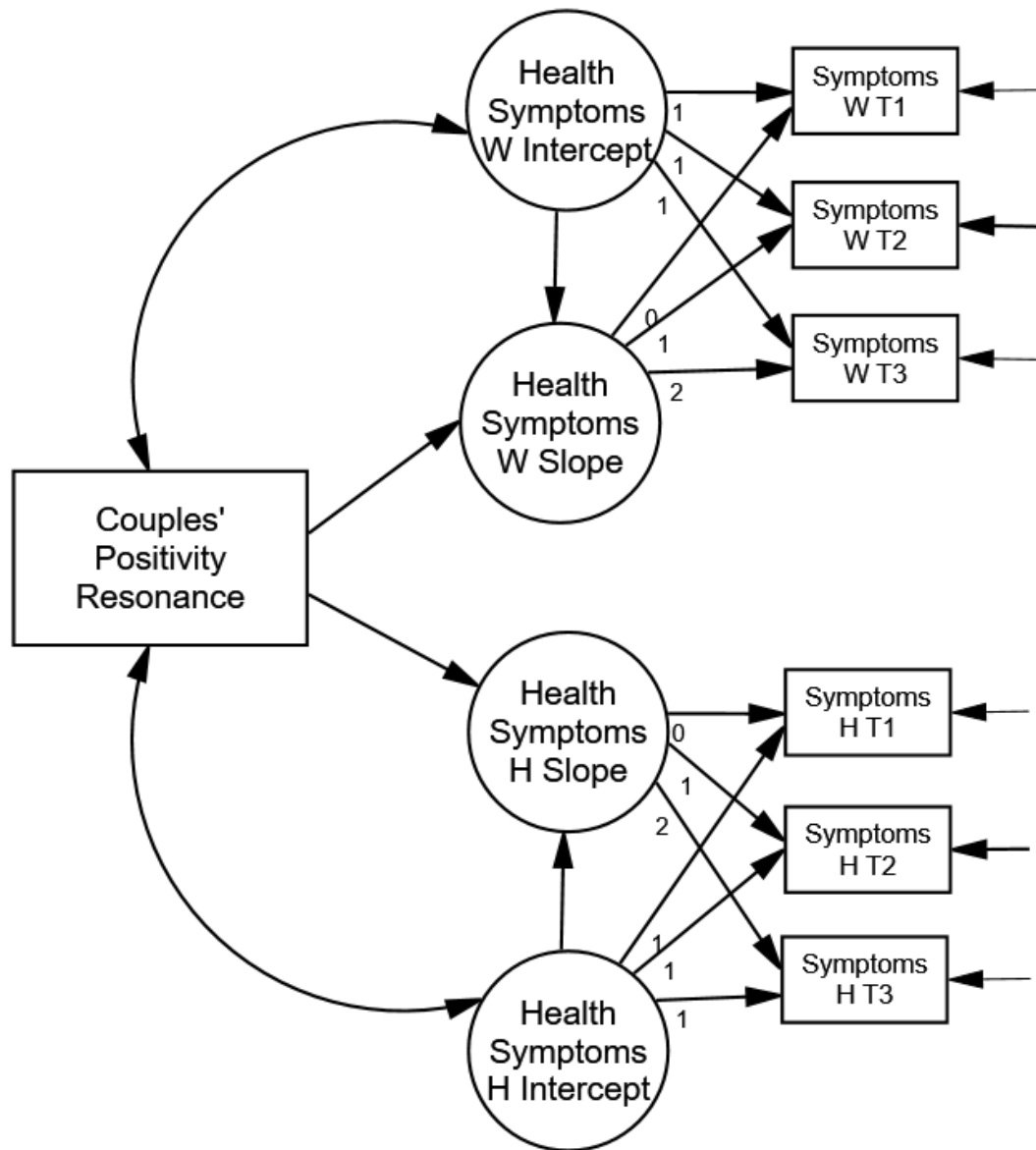
541 wives (separately) before constructing a dyadic LGM. To verify whether health trajectories
542 followed a linear pattern of change, we compared the dyadic LGM to a dyadic no-growth model.

543 *Longitudinal Health Predictions*

544 We used LGMs to examine how couples' factor-score latent positivity resonance at T1
545 predicted changes in both spouses' health symptoms over the ensuing 13 years (T1-T3). We
546 constructed a dyadic linear LGM with both wives' and husbands' health symptoms that included:
547 (a) intercepts (loadings of 1, 1, 1; indicating baseline levels of health symptoms at T1) and slopes
548 (loadings of 0, 1, 2; indicating trajectories of health symptoms from T1 to T3) for both wives and
549 husbands; (b) latent slopes regressed onto factor-score latent positivity resonance at T1; (c)
550 correlations between wives' and husbands' latent intercepts and factor-score latent positivity
551 resonance at T1; and (d) residual correlations within and across spouses' latent intercepts and
552 slopes (to account for the shared variance between wives' and husbands' health symptoms). To
553 test our hypotheses, we examined couples' factor-score latent positivity resonance predicting
554 wives' and husbands' health symptoms slopes, controlling for each spouse's own health
555 symptom intercept (e.g., in the regression with factor-score latent positivity resonance
556 predicting wives' slope, wives' intercept was included as a covariate). **Figure 1** depicts the
557 conceptual dyadic LGM.

558 **Figure 1**

559 *Positivity Resonance and Health Symptoms: Conceptual Dyadic Latent Growth Curve Model*



560

561 *Note.* Cross-spouse correlations between health symptoms intercepts and slope residuals as well
 562 as cross-spouse paths between health symptoms intercepts and slopes were also modeled but are
 563 omitted here for sake of clarity. Couples' positivity resonance was modeled as an observed
 564 variable, using factor scores to represent the latent construct that emerged from confirmatory
 565 factor analysis. W = Wives. H = Husbands. T1: 1989/90. T2: 1995/96. T3: 2001/02.

566 **Gender Differences.** To evaluate whether associations between couples' positivity
567 resonance and individuals' health trajectories differed as a function of gender, we fit another
568 dyadic LGM and constrained the effects of factor-score positivity resonance on health
569 symptoms slopes and the correlations between positivity resonance and health symptoms
570 intercepts to be equal across wives and husbands. We used a chi-square likelihood-ratio test to
571 compare the fit of the model with equality constraints to the initial dyadic LGM in which
572 associations with positivity resonance were estimated freely (Jöreskog, 1971).

573 **Covariates.** Given well-established associations of socioeconomic status (Adler &
574 Stewart, 2010) and health-related behaviors (McGinnis et al., 2002) with emotion and health, we
575 sought to examine prospective associations between positivity resonance at T1 and changes in
576 health symptoms over time by controlling for these potentially confounding influences.
577 Consistent with our prior work (e.g., Haase et al., 2016), analyses controlled for
578 sociodemographic characteristics (i.e., age, income, and education) and health-related behaviors
579 (i.e., a composite of smoking, alcohol consumption, caffeine consumption, and lack of physical
580 exercise) measured at T1 in the dyadic LGM. We also controlled for individually experienced
581 positive affect to evaluate the relative influence of dyad-level positivity, versus individual-level
582 positivity. These variables were included in the regressions with factor-score positivity
583 resonance predicting latent slopes, and we allowed for correlations between all covariates and
584 latent intercepts. Next, to investigate the added value of couples' positivity resonance beyond
585 self-reported marital satisfaction (which has already been linked with each of the dyad-level
586 indicators of positivity resonance; Brown et al., 2021; Chen et al., 2020; Otero et al., 2019), we
587 conducted additional LGM analyses following the same procedure as above, including wives'
588 and husbands' marital satisfaction at T1 as independent variables in the corresponding regression

589 analyses predicting latent slopes, and allowing for them to correlate with each other, with all
590 covariates, and with the latent intercepts.

591 **BIPR.** Finally, to explore whether the holistic behavioral measure, BIPR, would make
592 similar predictions for health trajectories to those made with the positivity resonance latent
593 variable (indexed by factor scores) we repeated all longitudinal health analyses with BIPR
594 (instead of factor-scored latent positivity resonance) as the independent variable of interest.

595 *Mortality Predictions*

596 We used Cox proportional hazard models to estimate the hazard ratios (HRs) and survival
597 curves for mortality (Cox, 1972). Specifically, we used shared frailty models, which incorporate
598 random effects to account for clustering of individuals within couples (Balan & Putter, 2020).
599 The shared frailty terms were assumed to have a log-normal distribution. Mortality analyses
600 proceeded in five steps. First, we assessed the proportional hazards assumption for all variables,
601 which assumes that the log hazard is a linear, time-invariant (parametric) function of the
602 predictors. In other words, it assumes the relative hazard remains constant over time for different
603 levels of each independent variable (Therneau & Grambsch, 2000). We included a time
604 interaction term for variables that violated this assumption (i.e., their effects on the HRs varied
605 over time) in all subsequent models, using the time-transform functionality of *coxph* in the
606 *survival* package (Therneau, 2020). Second, we tested whether factor-scored latent positivity
607 resonance predicted mortality. Third, we tested whether gender moderated any observed effect of
608 positivity resonance on mortality. Fourth, we examined whether factor-scored latent positivity
609 resonance predicted mortality, independent of sociodemographic (i.e., age, gender, income,
610 education), health (i.e., total health symptoms, health-related behaviors), affective (i.e.,
611 individually experienced positive affect), and relational (i.e., marital satisfaction) covariates.

612 Couples missing data for income ($n = 1$ couple) and individually experienced positive affect ($n =$
613 7 couples) were excluded from this step of analysis. Data for all other variables were complete.
614 Finally, we again tested whether BIPR would make similar predictions for mortality to those
615 made with factor-score latent positivity resonance by repeating analyses with BIPR as the
616 independent variable.

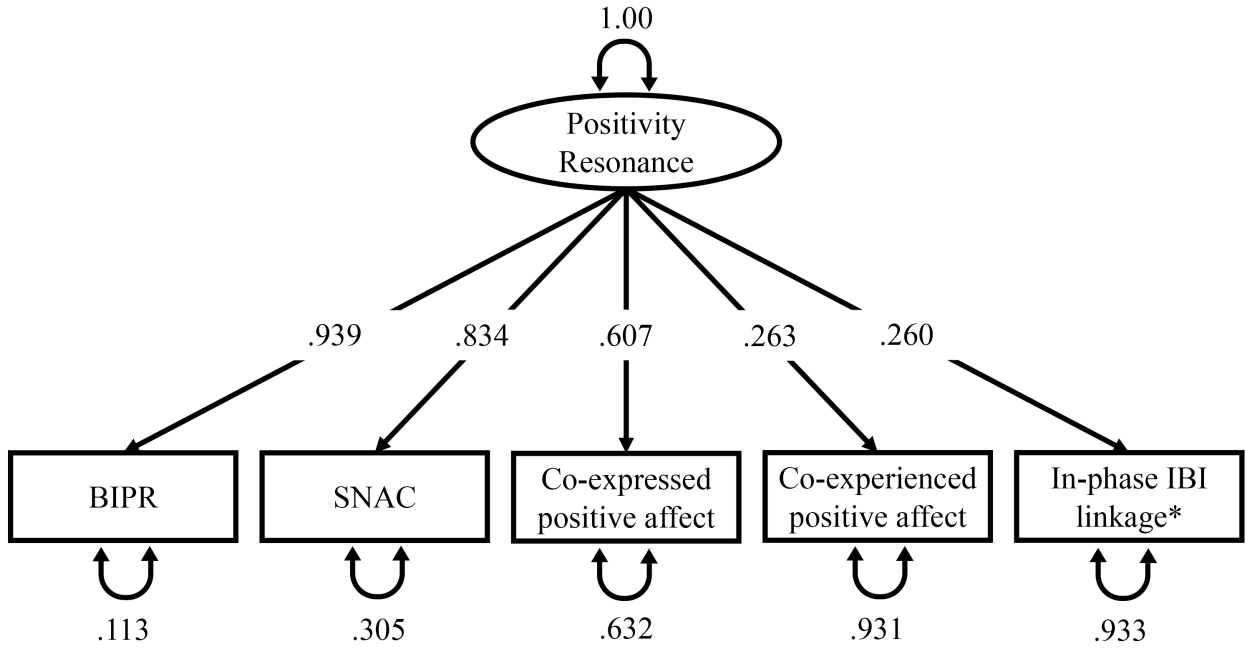
617 **Results**

618 **Preliminary Analyses: Construct Assessment**

619 *Measurement Model of Positivity Resonance (Hypothesis 1)*

620 We used CFA to test a measurement model of couples' positivity resonance, modeled as
621 a single latent factor indicated by BIPR, SNAC, co-expressed positive affect, co-experienced
622 positive affect, and in-phase IBI linkage during moments of co-expressed positive affect.
623 Supporting Hypothesis 1, the CFA for this model indicated excellent fit, $\chi^2(5) = 7.734$; $p = .172$;
624 CFI = .987; SRMR = .036. We found that all five measured indicators of positivity resonance
625 loaded significantly onto the latent variable (all $ps < .05$), with BIPR showing the highest loading
626 and co-experienced positive affect and in-phase IBI linkage showing the lowest loadings.
627 Because all loadings were significant, we did not exclude any indicators of positivity resonance
628 from the latent factor. Standardized factor loadings and residual variances are presented in
629 **Figure 2.**

630 **Figure 2**
631 *Confirmatory Factor Analysis of Positivity Resonance*



632
633 *Note.* All factor loadings were significant ($p < .05$). Curved arrows indicate standardized residual
634 variances. *during co-expressed positive affect. BIPR = Behavioral Indicators of Positivity
635 Resonance; SNAC = Synchronized Nonverbal Affiliation Cues; IBI = Inter-beat interval.

636 *Latent Growth Curve Modeling of Health Trajectories*

637 Separate linear LGMs of health symptoms showed good fit for wives and husbands, $ps \geq$
638 .366; CFI = 1.00; SRMR \leq .021. In the wives' model, the residual variance of wives' health
639 symptoms at T1 was negative and not significantly different from zero ($\delta = -.008, p = .945$),
640 thus, we fixed it to zero. A likelihood ratio test comparing an LGM with wives' T1 health
641 symptoms residual variance fixed to zero to the initial LGM showed that the models were not
642 significantly different ($\Delta\chi^2(1) = .004, p = .945$).

643 We proceeded to construct the dyadic LGM to model changes in both wives' and
644 husbands' health symptoms, which also showed good fit, $\chi^2(7) = 9.705; p = .206$; CFI = .993;
645 SRMR = .035. In the dyadic LGM, the only residual correlation that was significant was that
646 between wives' latent intercept and slope ($r = -.627, p = .027$). Husbands' latent intercept and
647 slope were not significantly correlated, nor were intercepts and slopes across spouses (all $ps >$
648 .05). Nonetheless, we included correlations between wives' and husbands' latent slopes and
649 intercepts to account for shared variance between wives' and husbands' health symptoms (akin
650 to modeling the shared frailty in survival analyses), following established procedures (Olsen &
651 Kenny, 2006).

652 We also compared the dyadic LGM to a dyadic no-growth model (Ferrer et al., 2004)
653 using a likelihood-ratio test and found that the dyadic linear LGM had significantly better model
654 fit ($\Delta\chi^2(9) = 18.471, p = .030$), thus, we continued to use the dyadic linear LGM in subsequent
655 analyses. The dyadic LGM showed that the mean health symptom score for wives at T1 was
656 18.95 with a positive but non-significant ($p = .455$) change across the ensuing 13 years (T1-T3),
657 whereas husbands' initial health symptom score at T1 was 13.44 with a positive slope that
658 approached statistical significance ($p = .062$). Therefore, the dyadic LGM fit the expected pattern

659 of change; the means of both wives' and husbands' latent slopes were positive, suggesting a
660 linear increase in health symptoms over time (i.e., health worsened over time).

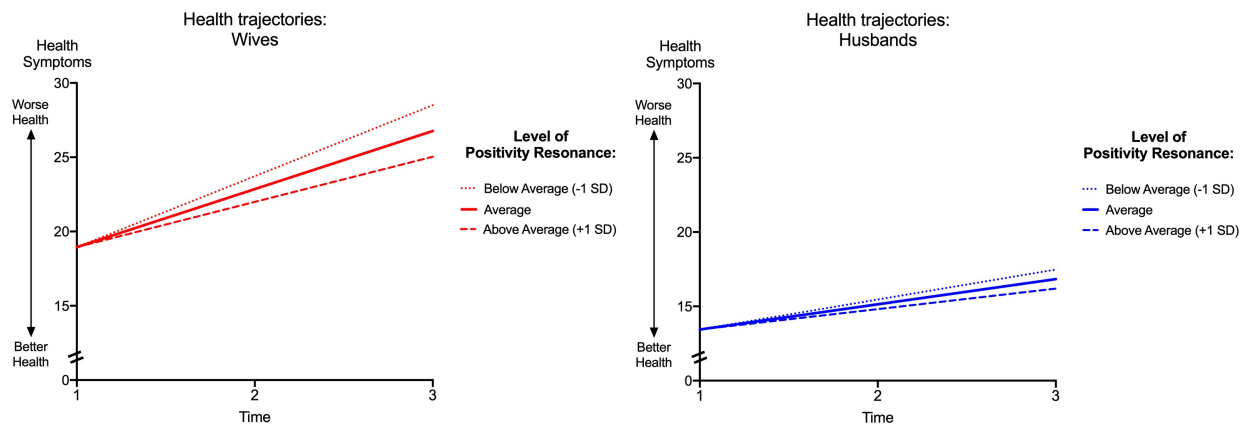
661 **Positivity Resonance and Longitudinal Health Trajectories (Hypothesis 2)**

662 We examined associations between couples' factor-scored positivity resonance at T1 and
663 changes in health symptoms from T1-T3 using a series of dyadic LGMs. All models showed
664 satisfactory fit ($p > .05$ for χ^2 tests; CFI values $> .95$, SRMR values $< .08$).

665 *Predicting Health Trajectories*

666 Couples' factor-scored latent positivity resonance at T1 was neither associated with
667 wives' health symptoms intercept, $p = .305$, nor husbands' health symptoms intercept, $p = .129$.
668 However, couples' factor-scored latent positivity resonance at T1 negatively predicted wives'
669 health symptoms slope ($\beta = -.192$, $SE(\beta) = .402$, $p = .028$), adjusting for wives' health symptoms
670 intercept. In other words, higher positivity resonance predicted less steep declines (i.e., better
671 trajectories) in health symptoms over time for wives. Additionally, wives' health symptoms at
672 T1 (i.e., health symptoms intercept) negatively predicted wives' health symptoms slope ($\beta = -$
673 $.634$, $SE(\beta) = .060$, $p = .002$). These findings were not found for husbands' health symptoms
674 slope ($\beta = -.110$, $SE(\beta) = .366$, $p = .369$). **Figure 3** shows the development of health symptoms
675 over 13 years for those with low versus high positivity resonance at T1.⁴

⁴ High values (> 3 standard deviations above the mean) exist at each timepoint. Given the nature of the data, we believe these are genuine scores that represent important sub-populations. For this reason, we chose to retain these values in our analyses. However, if we Winsorize these values (Tukey, 1962) by replacing them with the greatest observed value less than 3 standard deviations above the mean, we find the same pattern of significant results.

676 **Figure 3**677 *Wives' and Husbands' Health Trajectories Based on Levels of Positivity Resonance at Time 1*

678

679 *Note.* Lines depict estimated health trajectories from dyadic latent growth curve model with
 680 factor-score latent positivity resonance predicting health symptom slopes, controlling for health
 681 symptom intercepts. SD = standard deviation.

682

683 ***Gender Differences in Longitudinal Health Predictions***

684 To test whether the effects of positivity resonance on health trajectories were, in fact,
 685 statistically different for wives and husbands, we constructed a dyadic LGM using the same
 686 parameters as above, except we constrained the effects of couples' factor-score latent positivity
 687 resonance on health symptoms slopes to be equal across wives and husbands. We also
 688 constrained the correlations between positivity resonance and health symptoms intercepts to be
 689 equal across wives and husbands. In this model, couples' positivity resonance at T1 was not
 690 associated with health symptoms intercepts (across both wives and husbands), $p = .108$.
 691 However, couples' positivity resonance at T1 negatively predicted health symptoms slopes

692 across both wives ($\beta = -.129$) and husbands ($\beta = -.177$, $SE(\beta) = .285$, $p = .045$)⁵, when adjusting
693 for health symptoms intercepts. We then conducted a likelihood ratio test comparing the dyadic
694 LGM with imposed equality constraints to the initial dyadic LGM (where effects are estimated
695 freely across spouses) and found that the models were not significantly different ($\Delta\chi^2(2) = 1.275$,
696 $p = .529$). This null effect suggests that the effects of positivity resonance on health do not differ
697 significantly across genders. We proceeded to use the dyadic LGM with the aforementioned
698 equality constraints in subsequent analyses, given that it emerged as the more parsimonious
699 model.

700 *Robustness When Adjusting for Covariates*

701 **Sociodemographic Characteristics, Health-Related Behaviors, and Individually**

702 **Experienced Positive Affect.** Adjusting for individuals' age, income, education, health-related
703 behaviors, and individually experienced positive affect at T1, couples' positivity resonance at T1
704 was not associated with health symptoms intercepts, $p = .076$. When adjusting for these same
705 covariates as well as health symptoms intercepts, couples' factor-score latent positivity
706 resonance at T1 continued to negatively predict health symptoms slopes ($\beta = -.149$ for wives, $\beta =$
707 $-.155$ for husbands, $SE(\beta) = .282$, $p = .042$). Among the covariates, only husbands' health
708 symptoms intercept was associated with husbands' health symptoms slope ($\beta = -.383$, $SE(\beta) =$
709 $.069$, $p = .019$).

710 **Marital Satisfaction.** Adjusting for all the above covariates and individuals' marital
711 satisfaction at T1, couples' positivity resonance at T1 was not associated with health symptoms
712 intercepts, $p = .063$. When adjusting for marital satisfaction as well as health symptoms

⁵ The variances of wives' and husbands' health symptoms slopes are different, which leads to differences in the standardized regression weights. We imposed constraints on the raw regression weights because of their lack of dependence on variances. Standardized effects will differ across wives and husbands, but standard errors and p -values will be equal, in the models with equality constraints.

713 intercepts, couples' factor-score latent positivity resonance at T1 no longer significantly
714 predicted health symptoms slopes ($\beta = -.137$ for wives, $\beta = -.100$ for husbands, $SE(\beta) = .287$, $p =$
715 $.170$). Individuals' marital satisfaction also did not predict health symptoms slopes for wives nor
716 husbands, $ps > .133$); though it was associated with health symptoms intercepts for both wives (β
717 $= -.284$, $SE(\beta) = 1.198$, $p = .001$) and husbands ($\beta = -.208$, $SE(\beta) = 0.699$, $p = .010$).

718 ***BIPR and Longitudinal Health Trajectories***

719 In the CFA conducted in the preliminary analyses, BIPR (Otero et al., 2019) was highly
720 correlated with the latent positivity resonance factor and had the highest factor loading ($\lambda = .94$)
721 among all of the indicators. To evaluate whether BIPR by itself would have similar predictive
722 validity as did the latent factor (represented by factor scores), we repeated all longitudinal health
723 analyses, replacing factor-score latent positivity resonance with BIPR as the independent
724 variable. Re-running the above dyadic LGMs with BIPR, the overall pattern of significance
725 remained unchanged: BIPR-based positivity resonance at T1 continued to robustly predict the
726 development of health symptoms over 13 years ($\beta = -.129$ for wives, $\beta = -.178$ for husbands,
727 $SE(\beta) = .272$, $p = .043$). In sum, BIPR performed similarly to the latent factor of positivity
728 resonance in making longitudinal health predictions (i.e., standardized regression coefficients for
729 both measures were nearly equivalent, $\sim .20$). Full analyses using BIPR to predict longitudinal
730 health trajectories are presented in Online Supplemental Materials (see **Supplemental Results:**
731 **BIPR and Longitudinal Health Trajectories and Supplemental Figure S1**).

732 **Positivity Resonance and Longevity (Hypothesis 3)**

733 ***Proportional Hazards Assumption***

734 We assessed the proportional hazards assumption by fitting a Cox proportional hazard
735 model with all independent variables; obtaining the Schoenfeld residuals (i.e., the observed

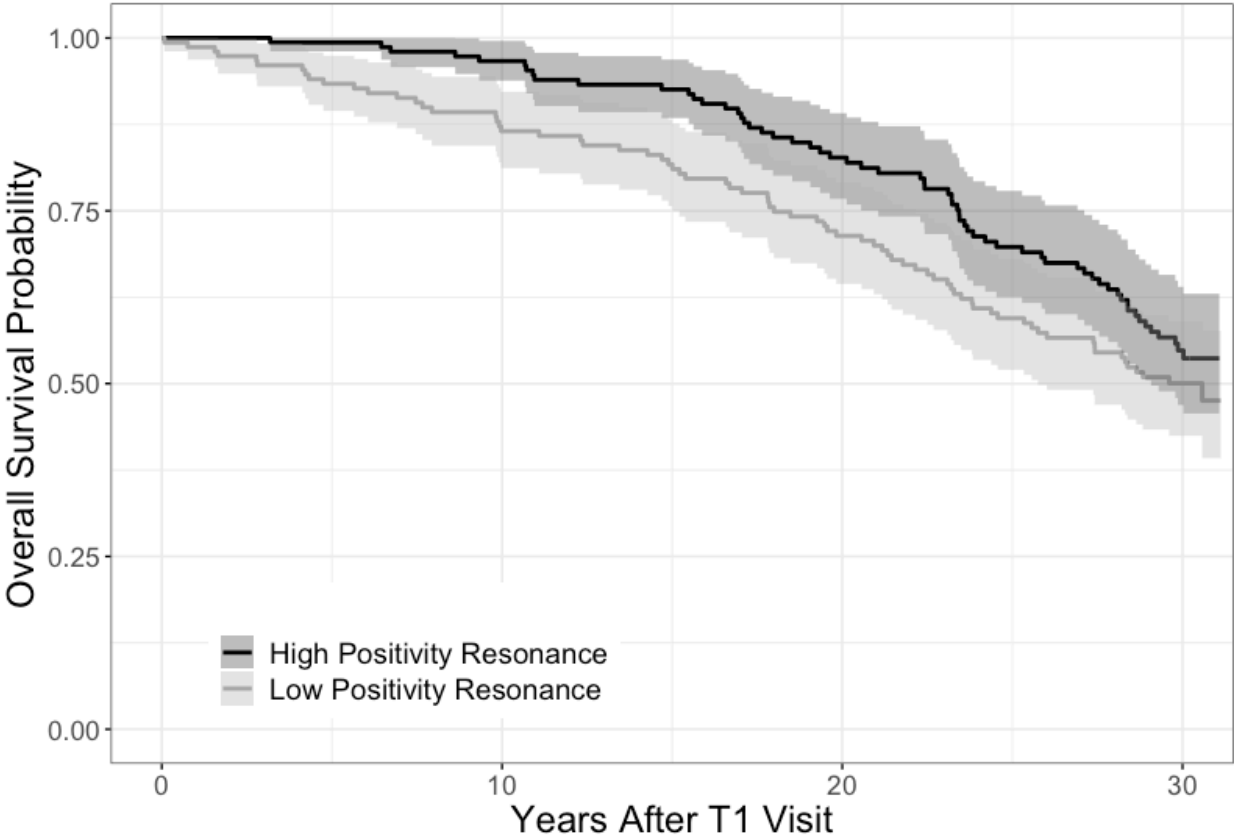
736 values of the predictors minus their predicted values at each event time; Schoenfeld, 1982); and
737 testing whether each variable exhibited a significant interaction with log-transformed time
738 (Grambsch & Therneau, 1994). Analyses revealed that the effects of couples' factor-score latent
739 positivity resonance ($\chi^2(0.90) = 6.61, p = .009$) and individuals' age $\chi^2(0.92) = 7.28, p = .006$) on
740 the Hazard Ratios (HRs) varied over time. A global test of non-proportionality showed that the
741 overall model did not violate the proportional hazards assumption ($\chi^2(20.96) = 16.27, p = .752$).

742 *Positivity Resonance Predicts Longevity*

743 We tested whether couples' factor-score latent positivity resonance (along with the
744 interaction of positivity resonance with time) predicted mortality. As depicted in **Table 3** (Model
745 1), greater positivity resonance predicted increased longevity such that there was a 78% decrease
746 in expected mortality for each standard deviation increase in couples' positivity resonance (see
747 **Figure 4** for survival curves). In other words, greater positivity resonance was associated with a
748 reduced risk of death. The interaction between positivity resonance and time also predicted
749 mortality, such that the strength of the effect of positivity resonance on mortality became weaker,
750 albeit slightly (i.e., the interaction effect HR = 1.00), over time.

751 **Figure 4**

752 *Survival Curves for Positivity Resonance and Mortality*



753
754 *Note.* Lines indicate estimated survival curves and shaded areas indicate 95% confidence
755 intervals around the associated survival curves. Couples' factor-scored latent positivity
756 resonance is depicted using a median split for display purposes only. T1 = Time 1.

757 ***Gender Differences in Mortality Predictions***

758 We tested whether gender moderates the association between factor-score latent
759 positivity resonance and mortality by including positivity resonance (along with a positivity
760 resonance by time interaction), gender, and an interaction term between positivity resonance and
761 gender in a model predicting mortality. Greater factor-score latent positivity resonance
762 continued to predict increased longevity (HR = 0.21, 95% CI [0.089, 0.481], $p < .001$), as did
763 female gender (HR = 0.56, 95% CI [0.385, 0.827], $p < .001$). The interaction term was not
764 significant, $p = .600$, providing additional evidence that the longitudinal health effects of
765 positivity resonance do not vary by gender. Therefore, we omitted the positivity resonance by
766 gender interaction terms in subsequent models.

767 ***Robustness When Adjusting for Covariates***

768 **Sociodemographic Characteristics, Health-Related Behaviors, and Individually**
769 **Experienced Positive Affect.** Next, we examined whether positivity resonance predicted
770 mortality, independent of age, gender, income, education, health symptoms, health-related
771 behaviors, and individually experienced positive affect. As depicted in **Table 3** (Model 2),
772 results revealed that greater factor-score latent positivity resonance remained a significant
773 predictor of increased longevity. Additional predictors of longevity included gender (being
774 female decreased the risk of expected mortality by 51%); household income (one standard
775 deviation increase in income decreased the risk of expected mortality by 22%); and total health
776 symptoms (one standard deviation increase in symptoms increased the risk of mortality by 42%).
777 The interaction between age and time was a significant predictor of mortality, such that the effect
778 of age on mortality increased over time. Taken together, these findings are consistent with well-
779 established risk factors for mortality from the literature, indicating that greater positivity

780 resonance, being female, and greater income may independently protect against the risk of death,
781 whereas greater age and greater health symptoms may independently increase the risk of death.

782 **Marital Satisfaction.** Adjusting for all the above covariates plus individuals' marital
783 satisfaction at T1, greater factor-score latent positivity resonance remained a significant
784 predictor of decreased mortality, as depicted in **Table 3** (Model 3). We also found that greater
785 marital satisfaction significantly predicted increased mortality (i.e., had a hazard ratio > 1);
786 however, we caution against interpreting that association by noting that the zero-order
787 relationship between marital satisfaction and mortality is not significant (see **Supplemental**
788 **Table S2** for zero-order associations between each covariate and mortality). Given that the
789 association between marital satisfaction and mortality emerges only when accounting for
790 positivity resonance, it is possible that this association is driven by the variation in marital
791 satisfaction that is unrelated to positivity resonance. It may be that some individuals whose
792 relationships are characterized by lower positivity resonance (and thus have increased risk of
793 mortality) overreported their marital satisfaction, perhaps to appear socially desirable. Noting
794 that positivity resonance was assessed objectively in this study, whereas marital satisfaction was
795 reported subjectively, is consistent with this speculation.

796 ***BIPR and Longevity***

797 Again, we repeated all mortality analyses, replacing factor-score latent positivity
798 resonance with BIPR as the independent variable. Re-running the above Cox proportional hazard
799 models with BIPR, the overall pattern of significance was consistent: BIPR at T1 continued to
800 robustly predict mortality (HR = 0.21, 95% CI [0.085, 0.519], $p < .001$), including when
801 adjusting for all covariates. The interaction between BIPR and time also significantly predicted
802 mortality, such that the effects of BIPR on mortality decreased slightly over time (HR = 1.00,

803 95% CI [1.00, 1.00], $p = .018$). See **Supplemental Results: BIPR and Longevity,**
804 **Supplemental Table S1, and Supplemental Figure S2.**

805 **Table 3**

806 *Cox Regression HRs of Positivity Resonance and Covariates Predicting Mortality*

	HRs and 95% CIs		
	Model 1	Model 2	Model 3
PosRes	0.22 [0.10, 0.51] ***	0.28 [0.12, 0.64] **	0.24 [0.10, 0.57] **
PosRes * time	1.00 [1.00, 1.00] **	1.00 [1.00, 1.00] *	1.00 [1.00, 1.00] *
Age	—	1.45 [0.80, 2.62]	1.40 [0.77, 2.52]
Age * time	—	1.00 [1.00, 1.00] **	1.00 [1.00, 1.00] **
Gender (1 = female)	—	0.50 [0.33, 0.75] ***	0.51 [0.34, 0.77] **
Household income	—	0.81 [0.65, 1.00]	0.78 [0.63, 0.97] *
Education	—	1.05 [0.84, 1.31]	1.10 [0.88, 1.37]
Health symptoms	—	1.32 [1.05, 1.65] *	1.41 [1.12, 1.78] **
Health-related behaviors	—	0.89 [0.73, 1.09]	0.92 [0.75, 1.13]
Individual ^a PA	—	1.00 [0.83, 1.21]	1.01 [0.83, 1.23]
Marital satisfaction	—	—	1.27 [1.01, 1.60] *

807 *Note.* HRs = hazard ratios. PosRes = factor-scored latent positivity resonance. PA = positive

808 affect. ^aIndividually experienced. * $p < .05$, ** $p < .01$, *** $p < .001$. An asterisk (*) in the variable

809 column indicates an interaction with time. A dash (—) indicates that the given variable was not

810 included within the model. All variables are at the level of the individual, with the exceptions of

811 factor-scored latent positivity resonance (and its interaction with time) and household income.

812 All variables were measured at the first timepoint.

Discussion

813
814 In the present study, we tested whether positivity resonance (measured both as a
815 multimodal latent factor and through a holistic behavioral coding system) predicts 13-year health
816 trajectories and longevity. A measurement model comprised of novel, dyad-level measures of
817 positivity resonance, each objectively assessed, had excellent fit, and thereby supported our first
818 hypothesis that the observed scores for these variables are influenced by an emergent, latent
819 construct (i.e., positivity resonance). Latent growth curve modeling showed some evidence that
820 both wives and husbands exhibited increases in health symptoms over time. Results also
821 supported our second hypothesis that greater positivity resonance (latent factor or BIPR) predicts
822 better health trajectories (i.e., fewer increases over time in health symptoms). This association
823 was initially found for wives only, although we did not find evidence that there was a statistically
824 significant difference in the effects of positivity resonance on health trajectories across wives and
825 husbands. When equality constraints were imposed, positivity resonance significantly predicted
826 health trajectories across both spouses, and this model emerged as more parsimonious than the
827 model in which the effects of positivity resonance were estimated freely. However, the
828 association between positivity resonance and health trajectories was not robust when accounting
829 for marital satisfaction, which was somewhat unsurprising given high multicollinearity among
830 positivity resonance and marital satisfaction (i.e., features of positivity resonance have been
831 consistently positively correlated with marital satisfaction in this sample; Brown et al., 2021;
832 Chen et al., 2020; Otero et al., 2019).

833 In another set of analyses, we found that greater positivity resonance (latent factor or
834 BIPR) predicted greater longevity (i.e., decreased risk of mortality), supporting our third
835 hypothesis. Again, gender did not moderate this association; and further, this association was

836 independent of self-reported marital satisfaction. Moreover, all associations (13-year health
837 trajectories and longevity) were robust when accounting for sociodemographic characteristics
838 (i.e., age, income, education), health-related behaviors (i.e., smoking, alcohol consumption,
839 caffeine consumption, and lack of physical exercise), and individually experienced positive
840 affect (e.g., the number of seconds in which wives reported feeling positive while husbands did
841 not). Results also indicated an interaction between positivity resonance and time, such that the
842 effects of positivity resonance on longevity were slightly attenuated over time. We speculate that
843 this time-related reduction in impact may reasonably reflect that other risk factors show time-
844 related increases in impact over time, like age, that may ultimately mitigate the long-term
845 protective effects of resilience factors like positivity resonance. Nonetheless, the robust
846 associations between positivity resonance and longitudinal health and longevity are particularly
847 striking given that these measures were drawn from one 15-minute conversation that occurred
848 over a decade (in the case of health symptoms) and up to three decades (in the case of longevity)
849 earlier. Taken together, these findings offer support for Positivity Resonance Theory, and
850 suggest that the novel group-level affective construct of positivity resonance may be an
851 important predictor of the long-term health and longevity. Akin to individuals' day-to-day health
852 habits of participating in physical exercise and eating nutritious food, their day-to-day habits of
853 cultivating positivity resonance with others may also function as positive health behaviors
854 (Fredrickson, 2016).

855 **Construct Validation**

856 Our results provide preliminary evidence validating the existence of a multimodal
857 positivity resonance construct that is indicated by dyad-level experiential, behavioral, and
858 physiological measures. The factor loadings from the CFA provide insight into the degree to

859 which the various measured indicators of positivity resonance are represented by the latent
860 factor. Given that BIPR is a holistic measure that encompasses multiple theorized components
861 (rather than one defining feature) of positivity resonance, it makes sense that BIPR has the
862 highest factor loading. Co-experienced positive affect, followed by in-phase IBI linkage during
863 co-expressed positive affect, showed the smallest (albeit significant) associations with the latent
864 factor, consistent with previous work showing that physiological responses tend to show less
865 coherence with other domains of emotional responses (i.e., subjective experience, behavior;
866 Mauss et al., 2005; Mauss et al., 2004). Nevertheless, all measures had significant factor
867 loadings, supporting the hypothesis that these key features – shared positive affect, caring
868 nonverbal synchrony, and biological synchrony – reflect a collective-level latent factor of
869 positivity resonance.

870 **Wives and Husbands**

871 Our initial test of Hypothesis 2 suggested gender-specific effects, in that couples'
872 positivity resonance predicted wives', but not husbands', health trajectories over 13 years.
873 Considering that women tend to have larger social networks (Phillipson, 1997) and receive more
874 social support (Turner & Marino, 1994; Umberson, 1992) than men, they likely have more social
875 interactions than do men. Further, women may also *cultivate* more positivity resonance in such
876 interactions, given that they tend to smile and laugh more than men (Bachorowski et al., 2001;
877 LaFrance et al., 2003), which, in turn, is known to elicit more positive affect in their interaction
878 partners (Bachorowski & Owren, 2001; Niedenthal et al., 2010). If so, longitudinal associations
879 between positivity resonance and health may be more likely for wives, who conceivably benefit
880 from a higher “dose” of positivity resonance, than for their husbands.

881 Another plausible explanation for this initial finding could be that men often underreport
882 their health symptoms, perhaps in part due to social roles that influence willingness to disclose
883 and communicate distress (Barsky et al., 2001; Kroenke & Spitzer, 1998). This gender-specific
884 tendency may be a potential source of bias in self-reports that may have artificially dampened the
885 mean health symptoms scores for the husbands (see **Figure 3**, which reveals husbands' self-
886 reported health symptoms to be significantly lower than that of wives across all timepoints),
887 which may have influenced our analysis of the association between positivity resonance and
888 health trajectories in men.

889 Nevertheless, when we fixed the effects of positivity resonance on health to be equal
890 across husbands and wives, we found that positivity resonance significantly predicted health
891 trajectories across both spouses, and further, this model emerged as the more parsimonious
892 option. Additionally, we found evidence that the effects of positivity resonance on health extend
893 beyond questionnaire data to a more objective, valid outcome – mortality. Indeed, our results
894 supported both of our hypotheses, that positivity resonance predicts longitudinal health
895 trajectories and longevity, across both wives and husbands. Therefore, we would expect to see
896 similar results across other types of relationships and genders beyond the heterosexual,
897 presumably cisgender cohort examined here. Future research is needed, however, to examine
898 positivity resonance in other types of dyads and relationship contexts.

899 **Theoretical and Practical Implications**

900 The present work is grounded in theories of affective science. Principally, this study is
901 motivated by the Positivity Resonance Theory of co-experienced positive affect, which proposes
902 that (a) shared positive affect, caring nonverbal synchrony, and biological synchrony reflect
903 moments of positivity resonance; and (b) together, these responses promote health and well-

904 being over time (Fredrickson, 2013, 2016). Positivity Resonance Theory builds on the idea of
905 emotion coherence – that emotions involve coordinated changes across behavioral, experiential,
906 and physiological response systems (Ekman, 1992; Levenson, 1994) – and extends it to dyad-
907 and group-level changes in emotion. Recent work with the present dataset shows that in-phase
908 physiological linkage is greatest during seconds in which both partners are simultaneously
909 expressing or experiencing positive affect (Chen et al., 2020), and additional work demonstrates
910 that greater coherence between subjective experience and physiology is associated with greater
911 well-being (Brown et al., 2020). Here, we show positive covariation of dyad-level emotional
912 responses within a broader temporal unit (i.e., the entire conversation). Therefore, this collection
913 of findings lends support to Positivity Resonance Theory and have the potential to support
914 emotion coherence theory. Notwithstanding the rich history of emotion coherence, we
915 acknowledge that the present analytic approach does not provide the same degree of temporal
916 precision (e.g., moment-by-moment) with which foundational studies in this area have been
917 conducted (e.g., Mauss et al., 2005; Rosenberg & Ekman, 1994).

918 Adding to the affective science methods literature, we offer additional support for the
919 holistic coding system, BIPR (Otero et al., 2019). BIPR’s high correlations with the latent
920 positivity resonance factor as well as with all of the observed indicators (see **Table 1**)
921 demonstrate the construct validity of this relatively new, dyad-level behavioral coding system
922 (Otero et al., 2019). Further, longitudinal associations with 13-year health trajectories and
923 longevity were nearly identical across the BIPR measure and the latent positivity resonance
924 factor. Evaluating positivity resonance through multiple, dyad-level behavioral, experiential, and
925 biological measures enabled us to affirm their theorized covariance through CFA. However,
926 future researchers seeking to measure high-quality moments of positive interpersonal connection

927 may prefer to measure BIPR alone, rather than the full latent factor, which would reduce
928 demands on time and resources while still making similar health predictions. Indeed, BIPR
929 coding is less time-consuming (e.g., only two weeks of training were needed, and two viewings
930 of 30-second video records; Otero et al., 2019) than many widely used behavioral coding
931 systems (e.g., SPAFF).

932 Affective scientists should also note that our findings indicate socially-shared positive
933 affect may be more powerful in promoting long-term health and longevity than is individually
934 experienced positive affect. At the same time, relationship scientists should note that social
935 relationships may be especially effective in promoting good health outcomes when shared
936 positive affect, nonverbal care, and synchrony are present. The presence of these features may be
937 particularly important for promoting health during moments of conflict (i.e., the context in which
938 they were measured in the present study), given that positive affect can “undo” the
939 cardiovascular activation produced by negative affect, an effect that has been shown both for
940 negative affect induced within tightly controlled laboratory studies (Fredrickson & Levenson,
941 1998; Tugade & Fredrickson, 2004) and for negative affect that arises during conflictual
942 conversations between husbands and wives (i.e., as in the present sample; Yuan et al., 2010).
943 This “undo” effect of positive affect likely also extends to co-experienced positive affect (c.f.
944 Prinzing et al., 2020), and may thus function to mitigate risks for cardiovascular disease.
945 Nevertheless, co-experienced positive affect has been found to predict marital satisfaction in
946 other conversational contexts (e.g., discussion of a pleasant topic; Brown et al., 2021); however,
947 additional work is needed to clarify whether this would extend to longitudinal health and
948 longevity.

949 On the one hand, social interactions that are marked by positivity resonance likely
950 support the formation and maintenance of close relationships, as is consistently evidenced by
951 positive associations between positivity resonance (holistic and individual measures) and marital
952 satisfaction (Brown et al., 2021; Chen et al., 2020; Lai et al., in prep; Otero et al., 2019). On the
953 other hand, pre-existing relationship satisfaction is likely to facilitate more frequent emergence
954 of positivity resonance. Associations between positivity resonance and marital satisfaction are
955 likely bidirectional. Although the association between positivity resonance and health trajectories
956 was not robust when accounting for marital satisfaction, the association with longevity was
957 found to be independent of self-reported marital satisfaction. It may be that associations with
958 health trajectories were relatively weaker due to attrition (i.e., couples with lower positivity
959 resonance at T1 had higher dropout rates at T3) or common method variance (i.e., health and
960 marital satisfaction were both measured via self-report questionnaire), whereas the association
961 with longevity was relatively stronger for the same reason (i.e., individuals with lower positivity
962 resonance were more likely to pass away, and there was no common method variance between
963 marital satisfaction and mortality). Nevertheless, in addition to promoting relationship
964 satisfaction, positivity resonance may also play a role in other relationship functions such as
965 partner responsiveness (a feature of positivity resonance; Reis, 2014), capitalization (Gable &
966 Reis, 2010), and expressed appreciation (Algoe et al., 2013), all of which may serve as
967 springboards for positivity resonance.

968 **Strengths and Limitations**

969 The present study had numerous methodological strengths, including (a) utilizing a
970 longitudinal dataset, enabling detection of health effects that develop over time; (b) measuring
971 positivity resonance through multiple, objective dyad-level methods, which are less vulnerable to

972 inflated associations with self-reported health through common method variance (c.f. Kelley,
973 1992); (c) examining longitudinal health through two domains, including 13-year health
974 trajectories and longevity over an even longer time interval; (d) adjusting for sociodemographic
975 characteristics and health-related behaviors known to influence health; (e) testing the predictive
976 validity of positivity resonance, independent of individually experienced positive affect and
977 marital satisfaction; and (f) demonstrating the predictive validity of BIPR, a parsimonious
978 measure of positivity resonance that can be readily implemented by future researchers.

979 There are also several limitations to note. Although in 1989 the CMI was considered
980 among the best health measures in the field (e.g., Aldwin et al., 2001) and there is a large body of
981 research supporting its validity (Weaver et al., 1980), our measure of health symptoms was
982 obtained via self-report rather than from more direct health measures (e.g., BMI, health care
983 utilization). An additional limitation includes the potential generalizability of the present study,
984 which utilized data from a racially and ethnically homogenous sample of heterosexual married
985 couples in the San Francisco Bay Area in the 1990s. It remains to be determined whether these
986 findings extend to other types of relationships (e.g., friends, homosexual couples, newlyweds,
987 parent-child dyads), other demographic groups, or to couples outside of this geographical region
988 or time period. It is also important to acknowledge that this sample consisted of people who lived
989 through times when gender roles were changing radically and that other generations might show
990 different findings related to gender. Additionally, this sample only included couples where
991 marital satisfaction scores of individual spouses fell within 20 points of one another, and thus
992 results may not generalize to couples who have larger discrepancies in their marital satisfaction
993 levels.

994 It also bears mentioning that while the CFA of positivity resonance includes measures of
995 all its defining features, the results do not preclude the possibility that another factor structure of
996 positivity resonance exists. That is, while the shared variance of these measures does reflect an
997 underlying, latent factor, there may be other ways of measuring positivity resonance (not
998 captured here) that could strengthen the assessment of the factor. Additionally, given that the
999 absence of positivity resonance does not imply the presence of negative affect (and vice versa),
1000 future researchers should evaluate whether shared negative affect, or negativity resonance,
1001 exhibits unique associations with health and longevity. Finally, our study was designed to
1002 evaluate the longitudinal associations between positivity resonance measured at baseline and
1003 changes in health symptoms over time. Future longitudinal studies should evaluate bidirectional
1004 associations to test the possibility of upward spirals between positivity resonance and health over
1005 time (see Fredrickson & Joiner, 2018).

1006 **Conclusion**

1007 The current study is the first comprehensive, multimodal assessment of positivity
1008 resonance at the dyadic level. Results lend support for our hypotheses that positivity resonance
1009 shows prospective associations with long-term health trajectories and longevity, which were
1010 observed to be independent of individually experienced positive affect. Conceptually, the high
1011 covariance observed among the defining features of positivity resonance offer further support for
1012 the Positivity Resonance Theory of co-experienced positive affect (Fredrickson, 2016).
1013 Methodologically, BIPR, the holistic behavioral coding measure, performed on par with the
1014 more comprehensive latent factor of positivity resonance in its health and longevity predictions,
1015 and may emerge as the most useful tool for researchers working in this area. The present findings
1016 also contribute to scientific understanding of interpersonal emotions and behaviors that lay the

1017 foundation for long-term health and longevity. Future research should explore specific biological
1018 and/or behavioral pathways through which positivity resonance is linked with health and
1019 longevity, as well as whether the findings extend to other types of dyadic relationships.
1020 Considering mounting evidence underscoring the importance of high-quality social connections
1021 in daily life, positivity resonance should be evaluated as a potential intervention target to
1022 determine if it can lead to improvements in health and well-being throughout society (c.f. Zhou
1023 et al., in press).

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Online Supplemental Materials

Previous Publications

Results from this NIA-supported study (R01 AG007476) have been reported elsewhere (Bloch et al., 2014; Brown et al., 2021; Carstensen et al., 1995; Chen et al., 2020; Haase et al., 2016; Haase et al., 2013; Holley et al., 2013; Levenson et al., 1993; Levenson et al., 1994; Otero et al., 2019; Pasupathi et al., 1999; Seider et al., 2009; Shiota & Levenson, 2007; Verstaen et al., 2020; Yuan et al., 2010) and will continue to support studies of interpersonal emotion in long-term marriages. The initial goal of the study was to recruit a sample of middle-aged and older couples who were representative of the racial, ethnic, economic, and religious makeup of the San Francisco Bay Area. The experimental sample was recruited in a three-stage process: 1) random telephone surveys conducted by a survey research company to assess population characteristics (i.e., marital satisfaction, age, ethnicity, religion, and socioeconomic status) of people living in the area, 2) initial screenings of prospective participants in which they completed a questionnaire packet, and 3) recruitment of couples from the pool of prospective participants who met the selection criteria established from the results of the random survey.

Prospective participants were recruited through advertisements in newspapers, radio, newsletters, bulletins, flyers, and placards. Nine hundred and sixty potential participants were screened to determine the final sample. Couples were recruited, on the basis of age and marital satisfaction, into one of four categories: (a) middle-aged satisfied, b) middle-aged dissatisfied, c) older adult satisfied, and d) older adult dissatisfied. Middle-aged couples had to be married at least 15 years, with wives between the ages of 40 and 50; older adult couples had to be married at least 35 years, with wives between the ages of 60 and 70. Marital satisfaction scores from the Marital Adjustment Test (Locke & Wallace, 1959) from the initial stage 1 telephone survey were used to establish the marital satisfaction selection criteria. Couples were required to live within a 10-mile radius of the University of California, Berkeley. Additional criteria reflecting the modal long-term marriages of the Bay Area (observed in the phone screening) included: (a) spouses were within 5 years of age of each other; (b) marital satisfaction scores fell within 20 points of each other; (c) the primary wage earner had not retired; (d) neither spouse met criteria for alcohol use disorder; and (e) English was the native language or language customarily spoken in the home. Researchers were generally successful in recruiting a sample that matched the demographic criteria established in the random telephone survey in terms of age, marital satisfaction, socioeconomic status, and religion; however, European Americans were oversampled, with a 17% greater representation of European Americans compared to the original target. One hundred and fifty-five of the couples were in first marriages, and childless couples were quite rare.

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Supplemental Results: BIPR and Longitudinal Health Trajectories

We examined associations between couples' BIPR scores at T1 and changes in health symptoms from T1-T3 using a series of dyadic LGMs. All models showed satisfactory fit ($ps > .05$ for χ^2 tests; CFI values $> .95$, SRMR values $< .08$).

Predicting Health Trajectories

Couples' BIPR at T1 was neither associated with wives' health symptoms intercept, $p = .430$ nor husbands' health symptoms intercept, $p = .154$. However, couples' BIPR at T1 negatively predicted wives' health symptoms slope ($\beta = -.201$, $SE(\beta) = .065$, $p = .027$), when adjusting for wives' health symptoms intercept. In other words, lower BIPR predicted greater increases in symptoms over time for wives. As in the latent variable analyses, this finding was not mirrored for husbands' ($\beta = -.108$, $SE(\beta) = .058$, $p = .369$). **Supplemental Figure S1** shows the development of health symptoms over 13 years for those with low versus high BIPR at T1.

Gender Differences in Longitudinal Health Predictions

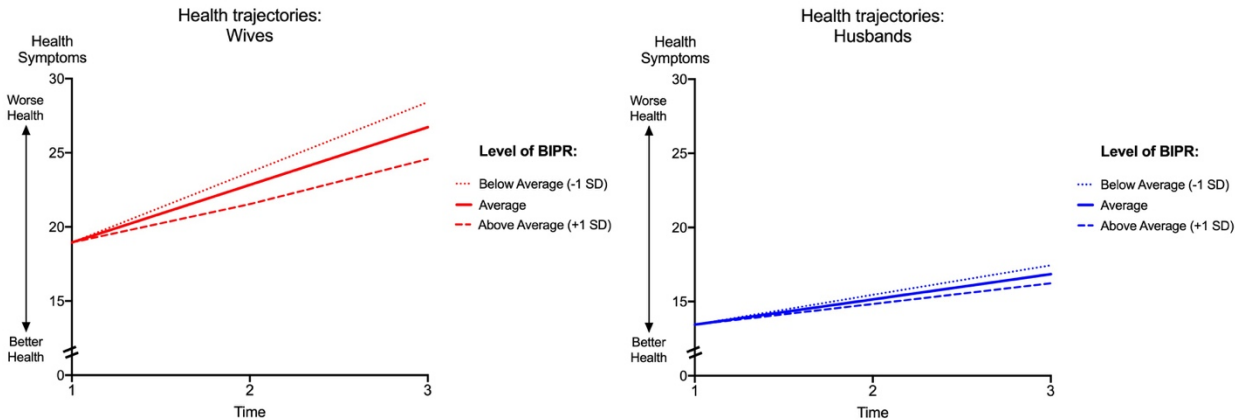
We constructed a dyadic LGM using the same parameters as above, except we constrained the effects of couples' BIPR on health symptoms slopes to be equal across wives and husbands; and we constrained the correlations between BIPR and health symptoms intercepts to be equal across wives and husbands. In this model, couples' BIPR at T1 was not associated with health symptoms intercepts (across both wives and husbands), $p = .163$. However, couples' BIPR at T1 negatively predicted health symptoms slopes across both wives and husbands ($\beta = -.129$ for wives, $\beta = -.178$ for husbands, $SE(\beta) = .272$, $p = .043$), when adjusting for health symptoms intercepts. We then conducted chi-square analyses comparing the dyadic LGM with equality constraints to the initial dyadic LGM (where effects are estimated freely across spouses) and found that the models were not significantly different ($\Delta\chi^2(2) = 1.274$, $p = .529$), suggesting that the effects of BIPR on health do not differ across genders. We proceeded to use the dyadic LGM with equality constraints in subsequent analyses, given that it is a more parsimonious model.

Robustness when Adjusting for Covariates

Sociodemographic Characteristics, Health-Related Behaviors, and Individually Experienced Positive Affect. Adjusting for individuals' age, income, education, health-related behaviors, and both husbands' and wives' individually experienced positive affect at T1, couples' BIPR at T1 was not associated with health symptoms intercepts, $p = .122$. When adjusting for these same covariates as well as health symptoms intercepts, couples' BIPR at T1 negatively predicted health symptoms slopes ($\beta = -.149$ for wives, $\beta = -.150$ for husbands, $SE(\beta) = .270$, $p = .043$).

Marital Satisfaction. Adjusting for all of the above covariates plus individuals' marital satisfaction at T1, couples' positivity resonance at T1 was not associated with health symptoms intercepts, $p = .111$. When adjusting marital satisfaction as well as health symptoms intercepts, couples' positivity resonance at T1 no longer significantly predicted health symptoms slopes ($\beta = -.141$ for wives, $\beta = -.102$ for husbands, $SE(\beta) = .274$, $p = .163$). Individuals' marital satisfaction also did not predict health symptoms slopes for wives nor husbands, $ps > .139$); though it was associated with health symptoms intercepts for both wives ($\beta = -.288$, $SE(\beta) = 1.212$, $p = .001$) and husbands ($\beta = -.207$, $SE(\beta) = 0.698$, $p = .010$). In summary, the same pattern of effects emerged with BIPR as those for the multimodal latent factor (reported in the main manuscript).

Supplemental Figure S1
Wives' and Husbands' Health Trajectories Based on Levels of BIPR at Time 1



Note. Lines depict estimated health trajectories from dyadic latent growth curve model with BIPR predicting health symptom slopes, controlling for health symptom intercepts. BIPR = Behavioral Indicators of Positivity Resonance; SD = standard deviation.

Supplemental Results: BIPR and Longevity

Proportional Hazards Assumption

We assessed the proportional hazards assumption by fitting a Cox proportional hazard model with all independent variables; obtaining the Schoenfeld residuals (i.e., the observed values of the predictors minus their predicted values at each event time; Schoenfeld, 1982); and testing whether each variable exhibits a significant interaction with log-transformed time (Grambsch & Therneau, 1994). Analyses revealed that the effects of couples' BIPR ($\chi^2(0.88) = 4.49, p = .028$) and individuals' age $\chi^2(0.90) = 8.93, p = .002$) on the HRs varied over time. A global test of non-proportionality showed that the overall model does not violate the proportional hazards assumption ($\chi^2(21.89) = 14.98, p = .859$).

BIPR Predicts Longevity

We tested whether couples' BIPR (along the interaction of BIPR with time) predicted mortality. As depicted in **Supplemental Table S1** (Model 1), greater BIPR predicted greater longevity, such that there is a 79% decrease in expected mortality for each standard deviation increase in couples' BIPR (see **Supplemental Figure S2** for survival curves). In other words, greater BIPR was associated with a reduced risk of death. The interaction between BIPR and time also predicted mortality, such that the effects of BIPR on mortality decreased slightly over time, albeit slightly (HR = 1.00, 95% CI [1.00, 1.00], $p = .018$).

Gender Differences in Mortality Predictions

We tested whether gender moderates the association between BIPR and mortality by including BIPR (along the interaction of BIPR with time), gender, and an interaction term between BIPR and gender in a model predicting mortality. BIPR continued to predict increased longevity (HR = 0.19, 95% CI [0.074, 0.491], $p < .001$), as did female gender (HR = 0.56, 95% CI [0.373, 0.836], $p < .001$). However, the interaction term was not significant, $p = .650$, providing additional evidence that the longitudinal health effects of BIPR do not vary by gender. Therefore, we omitted the BIPR by gender interaction in subsequent models.

Robustness When Adjusting for Covariates

Sociodemographic Characteristics, Health-Related Behaviors, and Individually Experienced Positive Affect. Next, we examined whether BIPR predicts mortality, independent of age, gender, income, education, health symptoms, health-related behaviors, and individually experienced positive affect. As depicted in **Supplemental Table S1** (Model 2), results revealed that greater BIPR remained a significant predictor of increased longevity. Additional predictors of longevity included gender (being female decreased the risk of expected mortality by 52%); household income (one standard deviation increase in income decreased the risk of expected mortality by 27%); and total health symptoms (one standard deviation increase in symptoms increased the risk of mortality by 43%). The interaction between age and time was also a significant predictor of mortality, such that the effect of age on mortality increased over time. Taken together, these findings indicate that greater BIPR, being female, greater income, and fewer health symptoms at T1 may independently protect against the risk of death.

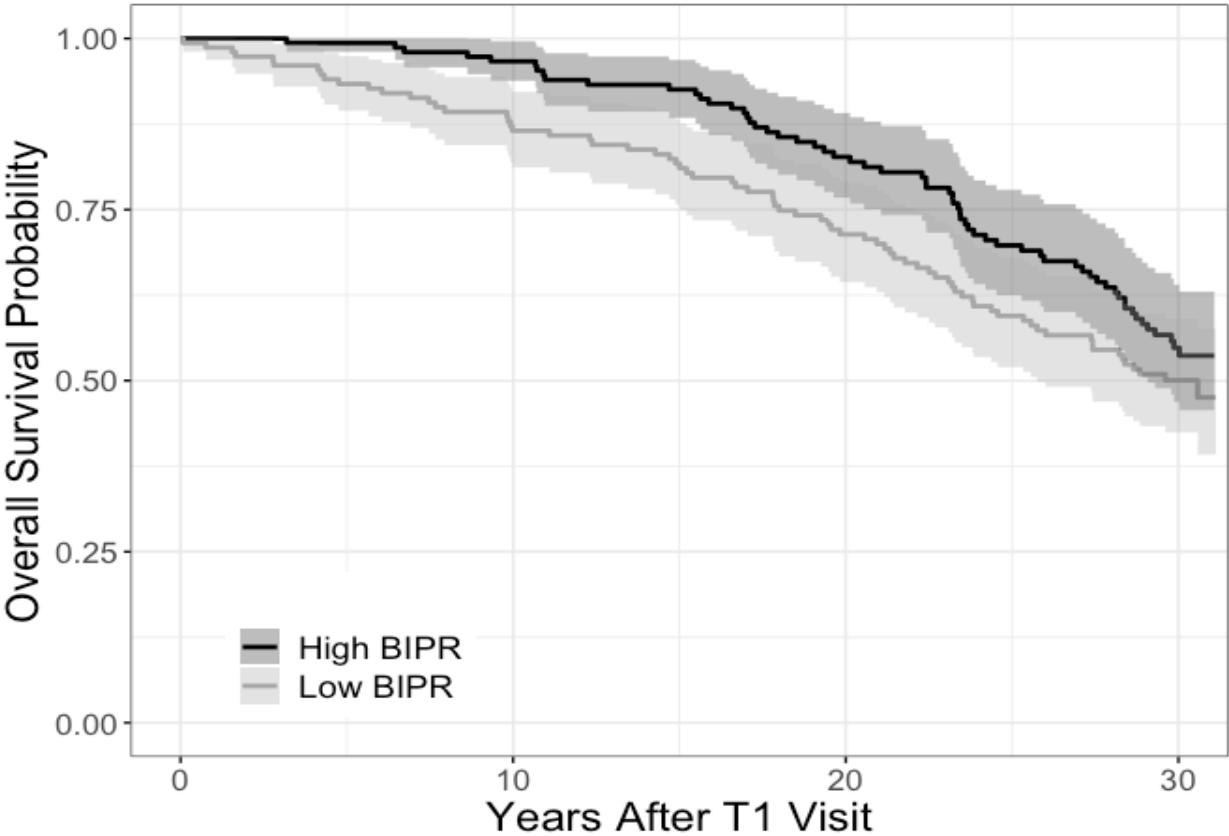
Marital Satisfaction. Adjusting for all the above covariates plus individuals' marital satisfaction at T1, greater BIPR remained a significant predictor of decreased mortality, as depicted in **Supplemental Table S1** (Model 3). Marital satisfaction did not significantly predict mortality.

Supplemental Table S1*Cox Regression HRs of BIPR and Covariates Predicting Mortality*

	HRs and 95% CIs		
	Model 1	Model 2	Model 3
BIPR	0.21 [0.08, 0.52] ***	0.27 [0.11, 0.67] **	0.24 [0.09, 0.63] **
BIPR * time	1.00 [1.00, 1.00] *	1.00 [1.00, 1.00] *	1.00 [1.00, 1.00] *
Age	—	1.46 [0.80, 2.67]	1.42 [0.78, 2.59]
Age * time	—	1.00 [1.00, 1.00] **	1.00 [1.00, 1.00] **
Gender (1 = female)	—	0.48 [0.31, 0.74] ***	0.48 [0.31, 0.75] ***
Household income	—	0.75 [0.59, 0.95] *	0.73 [0.58, 0.92] **
Education	—	1.01 [0.80, 1.28]	1.05 [0.84, 1.33]
Health symptoms	—	1.35 [1.07, 1.71] *	1.43 [1.12, 1.83] **
Health-related behaviors	—	0.90 [0.73, 1.10]	0.92 [0.74, 1.14]
Individual ^a PA	—	1.00 [0.82, 1.23]	1.02 [0.83, 1.25]
Marital satisfaction	—	—	1.21 [0.96, 1.53]

Note. HRs = hazard ratios. BIPR = Behavioral Indicators of Positivity Resonance. PA = positive affect. ^aIndividually experienced. * $p < .05$, ** $p < .01$, *** $p < .001$. An asterisk (*) in the variable column indicates an interaction with time. A dash (—) indicates that the given variable was not included within the model. All variables are at the level of the individual, with the exceptions of positivity resonance (and its interaction with time) and household income. All variables were measured at the first timepoint.

Supplemental Figure S2
Survival Curves for BIPR and Mortality



Note. Lines indicate estimated survival curves and shaded areas indicate 95% confidence intervals around the associated survival curves. BIPR = Behavioral Indicators of Positivity Resonance. Couples' BIPR is depicted using a median split for display purposes only. T1 = Time 1.

Supplemental Table S2

Cox Regression HRs of Zero-Order Associations Between Individual-Level Covariates and Mortality

	HRs and 95% CIs
Age	1.92 [1.13, 3.28] *
Age * time	1.00 [1.00, 1.00] *
Gender (1 = female)	0.57 [0.39, 0.82] **
Household income	0.86 [0.66, 1.11]
Education	1.05 [0.82, 1.33]
Health symptoms	1.25 [1.01, 1.54] *
Health-related behaviors	0.94 [0.76, 1.17]
Individual ^a PA	1.10 [0.90, 1.34]
Marital satisfaction	1.17 [0.91, 1.51]

Note. HRs = hazard ratios. PA = positive affect. ^aIndividually experienced. * $p < .05$, ** $p < .01$, *** $p < .001$.